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AN ACTUARIAL APPROACH

TO MOTOR INSURANCE RATING

Stewart Malcolm Coutts

 \mathbf{N}_{i}

A Thesis Submitted for the Degree

of Doctor of Philosophy

The City University Mathematics Department (Actuarial Science) London

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July 1983

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I would like to express my deep gratitude to both Professor Bernard Benjamin and Dr. Steven Haberman for their helpful supervision and constant encouragement throughout this research.

I am very grateful to the Insurance Company for providing the data and allowing me access to Company computer programs to perform the analysis. In particular I would like to thank Mr. Einion Holland and Mr. James Elmslie who encouraged all my research projects. Also Mr. Michael Brockman who provided all the computer output.

Special thanks are due to Professor Bobby Beard who taught me so much.

I would like to acknowledge numerous colleagues who over the years have contributed critical and positive advice, namely Dr. Lawrence Baxter, Professor Philip Dawid, Mr. Russell Devitt, Mr. Graham Masters and Mr. Graham Ross.

Finally I would like to thank Wyn Legon and Jan Stow who assisted in the final preparation of this manuscript.

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ABSTRACT

This thesis describes an actuarial structure for the practical analysis of motor insurance premium rating. An underlying theme emphasises that judgements are being made taking into account many factors e.g. economical, statistical and technical, therefore it is necessary to bring into the decision process a group of interested persons. In addition even though data are used to explain the proposed methods, it is the framework which is important and not the omission of some of the data e.g. important rating factors.

The basis for premium projecting is discussed together with a critical discussion of various measures of surplus. A new measure is developed referred to as 'proposed to existing' which measures the effect of premium adjustments after taking into account the portfolio distribution.

Another theme is to encourage a detailed within-portfolio analysis. An example, using data supplied by an Insurance Company helps to highlight the structure.

The analysis commences by sub-dividing the data into important underwriting rating factors. The claim experience is further divided by claim proportions and the three main types of claims cost: accidental damage, third party property damage and third party bodily injury. By sub-dividing the data into multiway cells both exposure and claim numbers become very small, hence statistical modelling is used to smooth the data and to reduce variation. A critical review of past models in respect of claim proportions and accidental damage costs is made. In addition a pragmatic approach to third party bodily injury is carried out. To obtain an office premium the modelled claim experience is combined with economic factors such as inflation and expenses.

Details of fitting the additive model by Orthogonal Weighted Least Squares is described. This converts the office premium into a 'points table'. An advantage of this 'points table' is that it can be used to compare various different sets of assumptions. A brief reference to the competitive market position is then made.

An analysis of surplus is developed together with a worked example, which highlights the importance of claim proportions and the level of claims cost.

Finally, the last chapter gives a summary of further research work which has been indicated as this thesis has developed.

KEY TO SYMBOLS AND ABBREVIATIONS

- AD Average accident damage cost per claim if settled immediately
- ACC Average claim cost
- ACF Average claim frequency
- AEP Average earned premium
- AOA Accident office association
- ASTIN Actuarial studies in non-life insurance
- AWP Average written premium
- BIA British Insurance Association
- CA Car age (factor)
- CO Cover (factor)
- COMP Comprehensive
- EAG Economic advisory group
- EP Earned premium
- EX vehicles exposed to risk
- FE Fixed expense
- GISG General insurance study group
- GIRO General insurance research organisation
- GLIM General linear interactive model
- IBNR Incurred but not reported

M Average miscellaneous cost per claim if settled immediately

- n Number of claims (associated with cell ijk)
- NON COMP Non comprehensive
- OP Office premium
- P Claim proportion (associated with cell ijk)
- PH Policyholder age (factor)
- RP Risk premium
- SB Standing business

- TPBI Average third party bodily injury cost per claim if settled immediately
- TPPD Average third party property damage cost per claim if settled immediately

US Underwriting surplus

VER Variable expense rates (including commission)

W Weight (associated with cell ijk)

VG Vehicle group (factor)

∧ Least squares estimate

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CHAPTER ONE

INTRODUCTION

The purpose of this thesis is to detail an actuarial approach to the technical aspects of motor insurance premium rating, where equal weights are given to both the practical and statistical elements of the problem. The methods described are applicable to the competitive UK motor insurance premium market. However, it is also contended that in countries where motor rating tariffs are in operation, the analyses proposed are still necessary for management to judge where in the portfolio the business is potentially unprofitable.

Only broad outlines of motor premium rating are available in the Institute of Actuaries literature (e.g. Beard (1964); Scurfield (1968); Johnson and Hey (1971); Benjamin (1977)). Elsewhere there is no shortage of theoretical papers on premium rating models (e.g. Pitkanen (1974). Bailey and Simon (1960) applied both a multiplicative and additive model to smooth claim ratios and applied their techniques to a set of Canadian motor insurance data. The first paper to analyse claim frequency and claims cost separately, together with expenses was that of Kahane and Levy (1975). The ASTIN Netherland Group (1982) presented a practical report on the premium motor structure in the Netherlands. The study was commissioned by the large insurance companies in the wake of the collapse of the tariff. The data analysed represented a significant portion of the Netherlands' private car population and must be

considered a major paper in reporting the results of a detailed premium analysis. The General Insurance Study Group (GISG), which is a forum in the UK, for actuaries involved in general insurance, presented four case studies on the practical aspect of premium rating namely; on household insurance Ajne, (1982); motor no claim discount systems, Christensen (1982); a motor insurance points table, Coutts (1982); reporting to the Insurance Commission in the USA, Grady (1982). Details of other important published work will be left to the relevant sections in the thesis.

A central theme of this thesis will be to emphasise a detailed within - portfolio analysis. The features taken into account will include:-

(a) some of the important underwriting rating factors

(b) statistical analysis of both claim frequency and claims cost

(c) expenses and inflation

These features will be combined to arrive at an office premium.

The statistical analysis will be applied to a set of data supplied by an Insurance Company. In some ways the data supplied does not take into account all the factors which ideally would be necessary to perform a full analysis. However, this thesis is attempting to give a framework for the whole analysis and therefore, these deficiencies will not invalidate the approach.

By performing this detailed analysis it is believed that:-

- the person responsible for the premium decision will be able to restrict his attention to the sensitive areas where judgements have to be exercised;

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- as the underlying assumptions are explicitly stated a monitoring process can be set up to establish if these assumptions are reasonable;
- any change in portfolio can explicitly be taken into account;
- this will enable management to establish where in the motor portfolio the business is either profitable or unprofitable.

The structure for this thesis is as follows:-

- Chapter 2 discusses the reason for a detailed data breakdown. In particular it is argued that the sub-division of the data, requires a statistical modelling approach, rather than using actual averages.
- Chapter 3 gives the general background to premium rating by highlighting the large time span between the data base information and the date when the claim will eventually be paid. Further it introduces the Group which is involved in the decision process. The chapter concludes with a critical review of the definition of surplus.

Chapter 4 gives the formula to be applied to arrive at an office premium.

Chapter 5 briefly defines the data which the Insurance Company supplied.

- Chapter 6 highlights the two main levels of decision namely, overall and within portfolio. Then the chapter concludes with a discussion about the problems concerning the multiway table which produces unequal numbers of claims per cell i.e. an unbalanced design.
- Chapter 7 reviews past published work on claims frequency and accidental damage claims cost statistical models. An additive model is then applied, to the data, where the fitting is achieved by Orthogonal Weighted Least Squares. Overall and within portfolio results are fully discussed.
- Chapter 8 discusses bodily injury analysis, from an overall and within portfolio view. Then details of a pragmatic approach is formulated to arrive at the input values.
- Chapter 9 is concerned with the economic factors relevant to the model. It briefly discusses fixed and variable expenses. Then lists the assumptions concerning past and future inflation which are used in the premium formula.
- Chapter 10 by using the Insurance Company's data, an actual office premium by rating factor combinations is calculated.

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- Chapter 11 discusses the presentation of office premium rates and introduces the concept of a 'points table'. Then by applying an additive model and fitting by Orthogonal Weighted Least Squares a 'points table' is derived.
- Chapter 12 compares different sets of assumptions used in the premium bases.
- Chapter 13 briefly discusses marketing aspects of premium rating.
- Chapter 14 analysis of surplus is developed and an example is discussed.

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Chapter 15 summarises the results and discusses areas for future research.

CHAPTER TWO

DATA BREAKDOWN

2.1 Overview

The first question to consider is whether the premium rates are to be reviewed and adjusted in an overall fashion, e.g. adding 10% over the whole portfolio, or whether to apply selective increases to different sections of the portfolio. It is argued that with the use of computers, a selective breakdown of important underwriting factors should be undertaken, which by aggregating the results can automatically give the overall level of premium adjustment needed. If, however, the system of analysis is only geared to the overall review, it is much harder to obtain information about the selective parts of the portfolio.

If the data are sub-divided by underwriting rating factors, this will lead to figures which are small both in exposure and claim numbers, so that simple averages will be suspect. Hence, it is suggested that simple statistical modelling be preferred. This has the following advantages:-

1. Extension of actuarial principles to small data-bases so that the portfolios of small Insurance Companies can and should be analysed. This statement is strongly worded, but it is believed that it can be accomplished since the theory of statistics in the past 10 years has made 'space age' progress

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in the analysis of small databases in far more critical and sensitive areas than motor insurance, namely, medical and demographic statistics, Little (1978). This formal statistical approach to the problem of small data-bases can be used to offset uninformed comments such as 'the data are too scanty to support any meaningful analyses'.

- 2. When the data are analysed in sufficient detail, then the effects of portfolio changes are reduced and judgements on these effects can be made with confidence.
- 3. A detailed analysis reveals that the process of premium rating involves many different assumptions. Changes in some of these assumptions can affect the premium rates significantly, for example different bodily injury assumptions (Chapter 12).
- 4. The establishment of statistical structures, however simple, provides 'bench marks' that can be used as a basis to monitor actual results as they emerge (Chapter 14).

2.2 Advantages and Disadvantages of an Analysis

Despite these points, it has been argued by non-actuaries that any actuarial input which might alter rates within the motor portfolio, is, practically irrelevant, when compared with overall marketing considerations or in countries where the motor rates are fixed by a tariff. Hence, the statistical process is considered a

mere theoretical exercise, the cost of which, given the personnel involved, is hard to justify. In addition, at the 1982 actuarial GISC Conference, a number of delegates supported this view. Their arguments were that an experienced actuary did not need to perform any detailed premium analysis, since the actuary should be aware of the premium situation and adjust the rates accordingly. It was also argued that past analyses should be a sufficient basis for changes in premium, if selected data were collected in order to monitor the process.

It is suggested that all these remarks are half-truths. Certainly, actuarial rates may differ considerably from market rates. However, it is the management's decision and they should be aware that the rates charged may in fact generate potential losses, the size of which should be quantified. If decisions are made without having all the relevant facts available, then this must be considered poor management. A topical example is that in the present market structure the underwriting rating factor 'car age' is generally ignored or given insufficient weight; however, from statistics available, it is evident that newer cars are being undercharged (Section 12.2). On the positive side, a detailed analysis may reveal unsuspected marketing opportunities within the present structure. In addition, in countries where there is fixed premium rating, set by a tariff, this may eventually break down, e.g. in 1982 in the Netherlands. Hence, it is advisable to be able to perform a detailed analysis to cover this eventuality.

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As for the argument that the experienced actuary has little need of analysis, it is accepted that judgements could often be made without any in-depth investigation, but how is this to be achieved? The existing body of knowledge has been accumulated largely by trial and error. In fact, however, the actuarial profession ought to strive to establish sound scientific procedures. It is contended that both the present UK and International actuarial literature does not give sufficient information for this purpose.

By appealing to past analyses, this tacitly assumes that premium structures are stable over time. It is accepted that frequent changes to the system may be undesirable, but regular analyses are necessary to verify assumptions made in the original calculations. In particular, if the data-base is small, this will necessitate regular checks to judge whether the past statistical inferences were reasonable. Thus in the UK, in the late 1970's, many companies reviewed their premium levels quarterly because of the rising rate of inflation. Furthermore, the underlying insurance risk pattern may also vary over time.

Finally on the financial side, cost and time are put forward as reasons for not performing regular analyses. However, with the advent of microprocessor technology, it is believed that the cost has been cut to a minimum and time reduced to an irrelevant factor.

In summary, a breakdown of the data to take into account rating factors is fundamental to establishing premium bases. To analyse data in this way, modern statistical techniques must be understood

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and employed regularly since the systems are not necessarily stable.As judgements are required to be made, it is necessary to set up monitoring systems.

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CHAPTER THREE

GENERAL BACKGROUND

3.1 Why Project?

Before discussing what background information is required in practice, to make judgements about premium levels, it would be useful to visualise the time span involved in the premium analysis. This is best explained by means of a simple example.

Consider a company that reviews its premium rates on 1st October 1980 and let us assume that these rates are expected to be in force for one year, Fig 3.1. New recommendations would have to be made in practice 3 months in advance. This time lag would be needed to alter computer output and prepare documentation for the rate book to be passed to the broker. The average policy will be effected halfway through the period over which the premium is expected to be in force, i.e. on 1st April 1981. This policy will be on risk for one year and, should it have a claim, the claim date will on average be 6 months after date on risk, i.e. 1st October 1981.

FIG 3.1

PROJECTION PROCESS

1980			1981			1983		
	JULY	OCTOBER	APRIL	OCTOBER	NNJ. <u>1</u>	OCTOBER		
	PREMIUM	PREMIUM	POLICY	CLAIM	VVV	TPBI		
	REVIEW	OPERATIVE	EFFECTED			PAID		

The material damage costs will be expected to be settled within 3 months from the date of accident. However, the third party bodily injury costs will take on average 2 years to settle. Prospectively, the average future time span is 3½ years from the date which the premium decision has to be made, i.e. by 1st July 1980. Hence, data concerning claims and expenses have to be projected to 1st October 1983 and beyond. This will be referred to as an 'averaging process'.

FIG 3.2

TOTAL TIME SPAN TO BE PROJECTED



Expense data will be based on information which is reasonably up to date at 1st July 1980 and is then projected to 1981. The claim data will be built up from claim numbers, material damage costs and third party bodily injury costs Fig 3.2. All but the latter costs will be based on recent data, say 1979/80. However, the most reliable third party bodily injury data are likely to be 5 years old, i.e. claims occurring in 1975. The reason why the average settlement figure of 2 years is not appropriate is because, in practice, the larger and proportionately more important claims take in excess of 5 years from date of accident to settlement. Hence, the total time span to be projected is on average $8\frac{1}{4}$ years. This large time span necessitates a number of subjective decisions, the main ones being:-

- Are third party claims occurring in 1975 relevant in respect of liability claims expected to occur in 1981 and to be settled on average in 1983 or later? The main problem lies in the settlement figures, as court judgements change with changing social conditions. This has been commonly referred to as 'judgement drift'. If the Company has an individual liability claim estimation process (referred to as a manual basis) then it is possible, though not necessarily reliable, to use later liability information based on recent manual estimates as a substitute for settled data.
- Whatever method is employed as the base for projecting claim costs and expenses, a view of past and future inflation has to be taken. In particular, for liability claims, a view of the rate of inflation to bring 1975 values up to 1980 is needed and thereafter a future rate to project these costs into 1983 or later.
- Finally, a view of future levels of claim frequency has to be decided. Factors including future weather conditions, petrol prices and speed restrictions have to be considered. It will be shown later that these aspects may have a relatively large effect on the profitability of results (Section 14.2).

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3.2 The Group

In a Company, it is desirable to make premium decisions within a Group. The decision-maker is the person who ultimately says what the premium rates are going to be; he would usually be a General Manager or the Motor Underwriter.

The rest of the Group would act as advisers to the decision-maker, supplying information on various aspects of the business. The size of the Group might range from one person for each main function, to, in a small company, as few as two, the decision-maker and the underwriter.

The following are the main aspects of the business which have to be considered:-

General underwriting principles, which vary from company to company and reflect, e.g. i) prior views on occupations or socio economic groups which the marketing should attempt to attract, e.g. teachers, civil servants; ii) wording in the policy conditions to take into account the introduction of a new rating factor such as protected No Claim Discount.

The overall market position of the Company compared to its competitors, with regard to growth and pricing. Within the Company, production summaries will be available showing lapse and new business figures. The marketing department will also be arguing for sales increases over selected parts of the portfolio, by trying to keep any rate increases to a minimum.

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Analyses from the claim personnel, who will report the latest manual estimates on present third party liability claims.

Statistical analyses will produce information for various members of the Group; concerning in particular, past claims frequency, claims cost and production results.

General economic factors will be used by the Group to project into the future past claim costs and expenses. This will involve, inter alia, a view as to the effect of the Government's current economic policy on inflation rates, for salaries and prices.

3.3 Definition of Surplus

Before the Group can judge the results of the latest premium analysis, it is necessary for them to define profitability. In the narrow sense, the profitability of an identical group of policies will depend on the balance between, on one hand, the premium receivable and on the other, claims and expenses incurred by that business. In general, premiums and associated expenses and some elements of claims cost will be known very quickly. However, as discussed above, the costs of third party liability claims may take some years to be accurately assessed. In practice, premium decisions have to be made before the ultimate claims costs are known, hence, the true profitability is not known when premium decisions are made. At any point of time, before the ultimate claim result is known, an estimate of the claim cost is made and when this is used, the term surplus will be employed instead of profit.

This definition of surplus is too general to be of use to management. It is necessary to sharpen the definition to provide both a specific measure of surplus together with its associated time period. Various definitions of these two latter concepts have been used by companies and herein lies the problem of establishing a universally acceptable monitoring procedure.

The above argument is known to actuaries. It is well documented in life assurance (Fisher and Young (1965)).

The two concepts, surplus and time periods, have to be related to achieve any meaning. Section 3.3.1 discusses surplus and Section 3.3.2 the time periods.

3.3.1 Measurement of Surplus

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In this section, several different definitions of surplus are discussed.

i) The claim or loss ratio is defined as:-

Estimated Total Cost of Claims Earned Premium

Where Earned Premium includes commission.

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This ratio shows the proportion of earned premium which is used to pay claims cost. Typically, the ratio will lie between 0.50 and 0.75. The main criticism is that this measure does not take into account expenses. Users say that comparison between companies requires a measure which is independent of expenses. However, for one's own company, the expense ratio is usually fixed and can be notionally added on at the end of the calculation, i.e. plus 0.25.

(ii) In the UK, in recent years, a measure which explicity takes into account expenses has been used, namely the operating ratio and is defined as:-

Claim Ratio + Premium Related Expenses + Other Expenses Written Premium Earned Premium

An example of premium related expenses is commission and of other expenses, salaries. For a discussion of this ratio, see Scurfield (1968). Typically, the values will lie between 0.95 and 1.10. The main criticism is the use of different denominators in the measurement which affect the sensitivity of the ratio, e.g. if written premium changes at a different rate to earned premium. This will happen when either the size of portfolio changes quickly or if there is a sudden increase in premium rates. This has the effect that underwriting surplus (see below) may not imply the same result as the operating ratio.

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(iii) Sometimes it is useful to look at the actual surplus rather than at a ratio. Then the underwriting surplus (or more usually known as underwriting profit) is defined as:-

earned premiums - estimated claims cost - expenses.

This measure will give the surplus in absolute terms and is usually used retrospectively. If used prospectively, it is very sensitive to portfolio change and requires some prediction as to the levels of classes of policies.

(iv) The following measure is not quoted in the literature. It compares a weighted average of two sets of alternative premiums, where e.g. one set could be based on the existing premium rates and the other on a new premium basis taking into account projected claim frequency, claims cost and expenses. The weights can either be the present or projected standing business. This measure will be referred to as the 'proposed to existing' basis.

 \sum Standing Business x Proposed Premiums \sum Standing Business x Existing Premiums Where the summation is over all policies

As this thesis is concerned with premium rates, the final measure 'proposed to existing' will be used to measure the effect of the office premium structure on the existing rates, and to compare the office premium to the fitted 'points table' (section 11.6.7).

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It is noticeable that none of the definitions of surplus take into account explicity, investment income. This omission will be briefly discussed later when the input assumptions are detailed. However, it is worth pointing out that, in general, the market ignores investment income for motor premium purposes. However, an implicit measure is used, where the assumed investment income expressed as a percentage of premium is added to the operating ratio.

For example, assume that the operating ratio = 1.03 and investment income is 10% of premium income. Then the trading result (underwriting surplus plus investment income) will give approximately a 7% return on premium income and the decision maker then has to decide if this is reasonable.

One reason why explicit account of investment income is not taken into account, is the difficulty of deciding what the future return on premiums will be.

3.3.2 Time Periods Associated with Surplus

As with the above measures, there are several associated accounting time periods used by companies to assess surplus. The problem is to relate premiums, claims cost and expenses to a defined period of risk.

- (a) The revenue account relates to a specific time period and credit is taken for earned premium, against which is set, claims and expenses received in that period. The problems which are discussed by Benjamin (1977) are briefly that it has no respect for the actual dates these values relate to, e.g. a claim which occurred in 1976 could have a significant adjustment in 1980 and this adjustment would be in the 1980 revenue account. The advantages with this method are that, it gives one figure for surplus quickly, it has been used by the insurance industry as the standard method of showing results and it seems to be reasonably well understood. The main disadvantage from a premium decision point of view is that it mixes up the experience of different premium decisions to a greater degree than the accident year method (see below). This means that any inadequacies of the present premiums will be hidden. The measures of surplus (i) to (iii) above can be used.
- (b) The year of accident basis considers all premiums, claims cost and expenses generated by policies at risk in a fixed period. The two main advantages of this method are that these data are required by the Department of Trade returns under the Insurance Companies Regulations (1981), hence data are being collected on an overall level and it shows claims cost development for different accident years.

However, like the revenue year, but to a lesser extent, this method mixes up the claim experiences relating to the different premium bases. The measures of surplus (i) to (iii) above can be used.

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(c) The policy year method considers the premium, claims cost and expenses generated from a cohort of policies which have been issued during a fixed period of time. Usually, the group of policies chosen relates to a specific premium basis. This method is satisfactory, since it allows the assumptions in the premium basis to be tested. However, to do this, it will be necessary to isolate the experience of this cohort separately. The main criticism is that the follow-up period would extend to two full calendar years and this might be considered too long a time to wait to test premium decisions. All measures of surplus can be used, but (iv) is preferred. A cohort analysis is not unique to general insurance, it is

widely used in demographic and medical studies.

3.3.3 Comparison of Time Periods

These three time periods are very different and they can effect the premium rates in terms of equity from the points of view of the Insured, the Insurance Company and the Supervisory body.

From the Insured point of view, the policy year method is the most equitable since it directly costs the premium prospectively. Hence, any shortfall in previous years premium bases will not be brought into the calculation. Interestingly, the Labour Government's Price Commission Regulation (for Insurance Companies) (1977) looked at the prospective cost of a cohort of policies and

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explicitly stated that any shortfall in previous premium rates could not be taken into account. The price control regulations were repealed in the late 1970's and since then, there has been no control on how to arrive at premium rates. Hence shortfalls of premium rates can be made up if the market allows. In addition, shortfalls are automatically made up under the year of accident method. Consider the following example.

Say, at 1st July 1980, we are interested in making an underwriting surplus equal to zero in 1981 (equivalent to making operating ratio = 1.00) by introducing a premium increase on 1st October 1980. To apply the definition of underwriting surplus it is necessary to estimate the claims cost for claims in 1981. This will take into account policies effected between 1st January 1980 and 1st October 1980, since they are at risk during the period between 1st January 1981 and 1st October 1981, but they will be subject to the present (i.e. pre-1st October 1980) rates. Hence, if the present rates are too low, this will effect underwriting surplus in 1981, as the earned premium generated by these policies will also be too low. Therefore, from the Insurers point of view, accident year is not very equitable. However, it is more equitable than the revenue year.

As far as the Company is concerned, the year of accident is a sensible method. The Company has to charge premiums, not only matching the prospective claims, but also taking into account the most recent past experience. The Company would argue that, since the rates under the policy year method are only estimates and everyone admits that these assumptions can be incorrect, is it so

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wrong to 'claw back' some previous underestimates in the recent past? (See below)

Also, if the previous inadequacies cannot be corrected explicitly within the rating structure, the Company will have to consider either raising more capital to finance this shortfall or making more conservative assumptions which would raise premium rates. This would also be in the interest of the Insured who would want the Company to safeguard itself against insolvency. The revenue account basis would be taking this argument to the extreme. It is believed that the year of accident is a reasonable compromise as far as the Company is concerned.

Finally, the supervisory body has to look at the situation from a solvency point of view, that is, it wants to be sure that the Company will pay retrospective and prospective claims (only IBNR) from the present and prospective value of assets.

However, the solvency regulations do not ask for any information concerning present premium bases and rely totally on past information to judge the solvency of that Company.

As a final comment, the above has only discussed the situation when previous premium rates are at best inadequate. Bohman (1979), in an interesting article, discussed methods where surplus is either repaid to the Insured by reducing premiums or accumulated to offset future inadequate rates. This is really a long term view of rates and is outside the scope of this thesis.

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As a summary of this chapter, it has been decided to consider the time period basis for calculation of premium rates to be related to a cohort of policies all affected during a particular calendar year. However, to obtain data on this basis would require an over-sophisticated data-base which is not available. Hence, a pragmatic approach has been taken, namely, as a compromise the averaging process outlined in Section 3.1 will be applied to data based on year of accident. The measure of surplus will be method (iv) in Section 3.3.1 namely a comparison of 'proposed to existing'.

CHAPTER FOUR

FORMULAE

It would be useful at this point to give the premium formula to be applied by combination of risk factors where appropriate.

The risk premium (RP), i.e. premium excluding expenses.

RP = fc

where $f = claim proportion = \frac{number of claims}{exposure}$

c = projected claims cost

=
$$AD(1+i)^{t}1 + TPPD(1+i)^{t}2 + TPBI(1+i)^{t}3 + M(1+i)^{t}4 \dots (4.1)$$

.

if settled immediately.

M = average miscellaneous cost per claim if settled immediately.

and

- ⁱ1 ^t2 ⁱ3 ⁱ4 are inflation rates for respective types of claim cost.
- t, t, t, t, t, t, t, are average settlement periods for respective types of claim cost.

The following comments can be made about equation (4.1):-

(a) An alternative suggested by Benktander (1982) is to replace $(1+i)^t$ by

$$\int_{0}^{t} e^{\int u_{dG}(u)} \stackrel{!}{=} e^{\int t}$$

Where the inflation rate is e^{5} , i.e. (5 - 1) p.a. and G(u) is unspecified with mean t the average time to settlement.

(b) Equation (4.1) divides claims costs, into AD, TPBI, TPPD and M, and so implicitly assumes that their respective proportion of the intimated claims remain stable.

The calculation of the office premium office OP, depends on how expenses are introduced. There are two main variants:-

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(i) Only fixed expenses:

$$OP = \frac{RP}{1-S} \qquad \dots (4.2)$$

i.e. a fixed percentage of premium

(ii) This method entails variable plus fixed expense.

$$OP = RP + (cc.f + \frac{t}{5} + r + ed) (1+j)^{\frac{1}{2}} \dots (4.3)$$

where cc = claims cost expenses

f = claim proportion

- nb = new business expense
- L = lapse expense
- t₅ = time period till lapse
- r = renewal expense
- ed = endorsement expense
- w = commission plus another expense related to
 premium
- j = inflation applicable to expenses

and the costs per unit are inflated to the relevant date of premium. A short discussion about expenses appears in Section 9.1.

Two omissions will be noticed from the office premium:-

Contingency loading (or solvency loading)

It is not certain whether companies explicitly take this factor into account. Although it has been omitted, its algebraic inclusion would be very easy. The problems are of estimating a value and its effect on the final rate. During the 1970's, when inflation was high there was a great deal of discussion on how to maintain solvency levels. One method looked at was to include in the premium basis a solvency loading. It is believed that in practice no company did this.

Investment Income

It is acknowledged that the market does not explicitly take this factor into account. Implicitly, income is taken into consideration when the whole motor account is scrutinised, in that the ultimate trading results can be compared with the underwriting results. Since equation (4.1) assumes the average date to settlement, i.e. payment; the adjustment to this equation to take into account investment income of k%pa is

$$AD \left(\frac{1+i}{1+k}\right)^{t} + TPPD \left(\frac{1+i}{1+k}\right)^{t} + TPBI \left(\frac{1+i}{1+k}\right)^{t} + M \left(\frac{1+i}{1+k}\right)^{t}$$

DATA

Private car motor insurance data were supplied by an Insurance Company. The Company also arranged the programming for all the grouped data by selected rating factors, showing vehicles exposed risk, numbers of claims, and the various average claims cost.

The list of information contained on the data file to arrive at these results are given in TABLE 5.1.

The exposure figures were obtained by the census method. Number of claims, AD, TPPD, TPBI and M, payment costs, were all related to the date of accident and the rating factors at that date.

The base data for exposure, claim numbers, AD, TPPD and M costs was 1 April 79 to 31 March 80. TPBI costs were for years 1972 to 1977.

It will be assumed that all the existing market underwriting factors will continue in use since it is unlikely that any underwriter would consider altering them without other companies following suit. In order to illustrate the principles under discussion, the rest of the thesis will be concerned with a worked example. The following rating factors will be used:-

TABLE 5.1

DATA CONTAINED ON COMPUTER FILE

Exposure Data	<u>Claims Data</u>
Policy Number	Claim Number
Cover	Policy Number
Policyholder Age	Cover
Car Age	Policyholder Age
Vehicle Group	Car Age
	Vehicle Group
Date of renewal	Date of Accident
	Payment of AD
•	Payment of TPPD

.

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×. .

Payment of TPBI

Outstanding amount

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•

(1)	Type of Cover	-	Comprehensive or					
			Non-Comprehensive					
(2)	Policyholder's Age (years)	-	17-20, 21-24, 25-29, 30-34,					
			35+					
(3)	Car Age (years)	_	0-3, 4-7, 8+					
(4)	Vehicle Group	_	A, B, C, D					

The vehicle group code A represents very small cars e.g. mini whilst D represents large or sports cars e.g. BMW or Morgan sports.

Two comments are necessary:-

(a) Some significant rating factors have been omitted from this list (e.g. district, no claim discount and use of car) since the data were not available. It is emphasised that these omissions may make the final results quoted in Chapters 11 and 12 not totally practical, in so much as the premium rate is not ready for quoting to a client. However, the main aim of this thesis is to present a framework for premium analysis rather than recommend a set of premiums for actual use by the company. Further as it is straight forward to extend the analysis to include these factors it is considered that these omissions do not invalidate the work.

(b) Following discussion with the underwriter, it may be possible to aggregate some of the detailed data into relevant groupings, in order to make the analysis more manageable. In our example, as an illustration, policyholder age 35+ is grouped instead of sub-dividing into 35-50, 50-60 and 60+. This sub-division is in fact quite popular in the market since it can be used to select retired people, or parents whose children have their own car.

CHAPTER SIX

THE GENERAL PRINCIPLES OF CLAIMS ANALYSIS

6.1 Objective

The principal objective of the analysis is to project past claims data for the relevant period. If the example in Section 3.1 is taken, then on 1st July 1980 a premium for the 1st October 1980 has to be decided, therefore all claims cost data will be projected to 15 August 1980. The premium rates will be in force for one year and the average date of claims arising will be 1st October 1981. Hence, all claim information has to be projected to 1st October 1981 and onwards. There are two levels of decision to make, namely,

(a) the overall levels of frequency and claims cost,

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and thereafter

(b) the within-portfolio levels, i.e. the relationship between rating factors.

The first stage will be dealt with by the Group with a minimum statistical analysis, but the ultimate decision requires a great deal of judgement. The second stage is basically where the statistical modelling and actuarial judgement become important.

6.2 Unbalanced Designs

Since the data-base of exposure and number of claims is relatively small the sub-division by rating factor, produces some cells with very small number of observed claims, making statistical modelling all the more necessary. The concepts of modelling are fundamental to a proper statistical analysis, Fisher (1946). In particular, modelling techniques are used to make inference about the structure and reduce variation. Dawid (1980) discussed the advantages of modelling for this type of data.

The main statistical problem is not the size of the data-base but the estimation of the parameters and the subsequent analysis of variance table for the multiway table which has an unequal number of observations (i.e. exposures and claims) per cell. This is also referred to as an unbalanced design. By inspecting TABLE 7.5 (to be discussed in the next chapter) which shows for the portfolio the actual exposure column (1) and claims column (3), the largest numbers occur for comprehensive cover for policyholder age 35+ and vehicle groups B and C. Whilst for non-comprehensive, small numbers occur for newer cars and all policyholder ages.

Bailey and Simon (1960), Seal (1968) and Johnson and Hey (1971), all acknowledged the lack of balance in the design and restricted their analyses to estimation of the parameters (without interaction terms). However, the advantage of extending the analysis to consider the relative importance of rating factors, was not discussed in depth. It is believed that part of the reason was the technical statistical difficulties.

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Francis (1973) bought to the notice of statisticians that many of the standard computer packages produced inconsistent results for the unbalanced design. The reasons are discussed fully in Aitkin (1978) - but the principal reason is the number of parameters included in the model. For the simple additive model with one observation (or equal number of observations) per cell, the sum of squares in the analysis of variance table are partitioned, that is for each parameter the explained variation will not change if an additional parameter is included, this is known as an orthogonal model. The unbalanced design is non-orthogonal and the explained variation in the analysis of variance table will alter if a parameter is included or excluded from the model. Nelder and Wedderburn (1972) developed the general linear model and included an analysis of variance which because it was generalised they called an analysis of deviance. The computer package called GLIM (1975) (General Linear Interactive Model) solved all the programming problems. However, to run successfully the GLIM package on the size of data included in this thesis a large amount of computer space is required (over 500K). The limitations and lack of access to a large computer was one of the reasons for the development of Orthogonal Weighted Least Square (Coutts (1975)).

GLIM was introduced to actuaries by Baxter Coutts and Ross (1980), however it has not gained any formal recognition. Little's (1978) work is important, since it explains in very simple terms the underlying principles of GLIM and leaves this reader convinced that is a solution for a large number of problems. Albrecht (1982) reviewed all the general linear model literature and is theoretically a very important paper.

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CHAPTER SEVEN

THE ANALYSIS OF CLAIM FREQUENCY AND MATERIAL DAMAGE CLAIMS COST

7.1 Introduction

This chapter describes the detailed analysis performed on both claims frequency and material damage claims cost. It also contains a critical review of past published work and concludes with the input values for the office premium calculation.

7.2 Claims Frequency: Overall Levels

Data for past years and quarters, showing actual claim proportions would be shown to the Group. From experience it is only necessary to decide on the overall levels for both the comprehensive and non-comprehensive sections of the portfolio. TABLE 7.1 shows the data available for comprehensive.

	TABLE 7.1									
	CLAIM FREQUENCY									
	OVERALL COMPREHENSIVE									
	Claim proportions per									
	1000 vehicles									
	Year of Accident Quarter of Accident Year									
		1	2	3	4					
	1977	142	125	115	152	138				
	1978	140	127	128	144	135				
	1979	184	130	131	154	150				
	1980	149	130*							
* includes an IBNR estimate										
where	Claim Proportion = _	Num	ber o	f cla	ims in	quarter				

Vehicles exposed to risk in quarter

Discussion on projecting the 1980 results into 1981 would be centred around such items as:-

- (i) weather conditions,
- (ii) petrol prices,
- (iii) road repairs,
 - (iv) general economic conditions, since these might affect the frequency with which policyholders have their cars serviced.

It is assumed that the Group decides to use the 1978 overall levels i.e. approximately 135 per 1,000 vehicles as the basis for projection for 1981. The data indicate that apart from the first quarter 1979, the years 1978 and 1979 were very similar, and that there is no reason to think that 1981 would be any different.

A similar table is prepared for non-comprehensive:-

TABLE 7.2									
CLAIM FREQUENCY -									
OVERALL NON COMPREHENSIVE									
Claim proportions per									
	10	00 vel	nic	les					
Year of Accident	Quart	er of	Aco	cident	<u>Year</u>				
	1	2	3	4					
1977	90	81	89	99	90				
1978	93	80	85	180					
1979	106	83	75	103	92				
1980	93	81*							
	* inc	ludes	an	IBNR	estimate				

For consistency with comprehensive the 1978 levels are assumed as the projection for 1981, i.e. approximately 89 per 1,000 vehicles. Finally the proportions of AD, TPPD, TPBI and zero claims were investigated for both comprehensive and non comprehensive. From the data, the assumption that these remained reasonably constant overtime were accepted.

7.3 Claims Frequency within Portfolio

7.3.1 Review

Statistical modelling techniques for claim frequency data are well documented, e.g. Ferrara (1971) and Bennett (1978) Many models have been used where the dependent variable was either the number of claims or claim proportions and the independent variables were the underwriter's rating factors. Interest has been centred on comparing additive and multiplicative models, and the different statistical procedures used to estimate the effect of the rating factor parameters.

Almer (1957) first suggested a multiplicative model, similar to

$$P_{ijk} = S R_i U_j V_k + E_{ijk} \dots (7.1)$$

where

^Pijk is the claim proportion for cell ijk

S is overall mean

Ν.

 R_i is the effect of rating factor R at level i U_j is the effect of rating factor U at level j V_k is the effect of rating factor V at level k E_{ijk} is the error term with mean zero and

variance Ø

Bailey and Simon (1960) investigated claim ratios by comparing both Almer's multiplicative model and an additive model, using an additive model similar to

$$P_{ijk} = S + R_i + U_j + V_k + e_{ijk}$$
 ... (7.2)

where e_{ijk} is the error term with zero mean and variance τ^2

The parameters were estimated by minimum chi-squared statistic.

Then followed a number of papers adapting the Bailey and Simon work on multiplicative models, e.g. Mehring (1964) Jung (1968) and Ajne (1974). It would appear that the multiplicative model become very popular in both the USA and Europe but not in the UK. Seal (1968) reintroduced interest in the additive model and used as the dependent variable a standard statistical transformation of log of the odds (or logit) namely

$$\log \frac{P_{ijk}}{1 - P_{ijk}} = S + R_{i} + U_{j} + V_{k} + E_{ijk}^{1} \qquad \dots (7.3)$$

The standard reasons for the transformation is that $\log \frac{P}{1-P}$ can be shown to be approximately normally distributed and the conditon that the estimated P lies between 0 and 1 is always satisfied, Cox (1970). It has been found in practice, that if there are a number of actual claims near to zero, then the fitted estimates for that cell could be negative, hence the advantage of the logit transformation. The fitting criterion was a weighted sum of squares where the weights were

 $\frac{P_{ijk}(1-P_{ijk})}{n_{ijk}}$

and n isk are number of vehicles exposed to risk in cell ijk. As the weights depended on the parameter estimates, the solution required an iterative procedure. Seal's work did not seem to make an impact on the UK actuarial profession. It seems that general insurance practical problems only came of age after the major paper by Johnson and Hey (1971). They introduced a similar model to Seal (1968), but replaced the weights by simply the number of vehicles exposed to risk, i.e. they assumed that P(1-P) was constant, for cells ijk and ignored the logit transformation. The fitted method was weighted least squares, i.e. minimise

$$\sum_{ijk}^{n} n_{ijk} \left(P_{ijk} - (S + R_{i} + U_{j} + V_{k} + ...) \right)^{2}$$
(7.4)

where is the least square estimate.

The analysis and results were based on a large data set and Johnson and Hey's conclusions were that, the main effects explained the underlying structure.

Johnson and Hey (1971) did not give the mathematical analysis, but left it to Grimes (1971). Grimes's paper was difficult to follow, but the underlying theory is explained in Coutts (1975). In the early 1970's, computer time was relatively expensive, hence Bailey and Simon's iterative procedures were not encouraged and the very large matrix inversion potentially required in weighted least squares was not discussed in

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actuarial literature. Johnson and Hey acknowledged these practical problems and limited their analysis to the main effects.

The computer size restrictions and the problems concerning the analysis variance table (section 6.2) prompted Coutts (1975) to developed an approximation to the standard weighted least squares analysis called Orthoganal Weighted Least Squares (OWLS). The approximation depended on the factorisation of the weights and was suggested by Please (1974). The method is briefly discussed in Appendix 1. Coutts (1975) showed that weighted least squares and OWLS with a logit transformation gave very similar results on a small data set. The main advantage of OWLS is that no computer inversion is necessary and a very large number of rating factors can easily be analysed and, in particular, the importance of the interaction terms can be investigated. (See Section 7.3.2).

A powerful statistical tool when fitting data to an additive model is the analysis of variance which has been mentioned in Chapter 6. However any inferences made, depends on the assumption of the model being satisfied. Coutts (1975) and Baxter and Coutts (1977) and Baxter, Coutts and Ross (1980) discussed the variation explained by the rating factors, and brought evidence to show that (on a data set supplied by the same Insurance Company that supplied data for this thesis) approximately 70% of the variation can be explained by the model of main effects. This was contrary to the result by Johnson and Hey (1971) who said that a very small amount of variation can be explained. The reason for this apparent difference of view, lies in the definition of variation being explained. In Baxter, Coutts and

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Ross (1980) the variation being measured is for a group of policies falling into cell ijk. Whilst in Johnson and Hey (1971) the variation being measured applies to an individual policyholder. (Johnson (1980)).

The major criticism of work published in the 1970's was discussed by Baxter, Coutts and Ross (1980), i.e.

"There is a tendency among many authors to advocate a particular model and proceed to estimate the parameters, often making no formal attempt to specify their assumptions concerning the error structure, and make no attempt to justify their choice in a statistical sense".

The paper went on to outline a statistical framework for modelling, namely: state the underlying assumptions, carry out the analysis and then attempt to verify the assumption by examination of the residuals. An analysis was performed on several data sets using GLIM. The paper demonstrated on relatively small data sets

- (a) that the multiplicative and additive models gave similar results (confirming Bailey and Simon (1960));
- (b) that for the error distribution Poisson (Jung 1968), binomial and normal assumption (Johnson and Hey (1971)) gave similar results;
- (c) that the OWLS gave similar results to the correct model even if the weights did not factorise. In particular the approximate analysis of variances table produced by OWLS gave similar results to the correct analysis of variances produced by GLIM. The analysis of variance table showed that approximately 70% of the variation was explained by the main effects;
- (d) That the standardised residual plots supported the model with main effects, hence the main effects were used to model claim proportions.

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7.3.2 The Analysis

In the spirit of the Baxter, Coutts and Ross (1980) paper and because the data analysed in this thesis is very similar, the following model was used to smooth the actual claim proportions

$$\log \frac{P_{ijk}}{1-P_{ijk}} = T + CO_{i} + PH_{j} + VG_{k} + CA_{l} + (In) + E_{ijkl}...(7.5)$$

where P = claim proportion for cell ijk

Т	= overall mean
CO_{i}	= rating factor cover (levels comprehensive and
	non-comprehensive
. PHj	= policyholders age (17-20, 21-24, 25-39, 30-34, 35 +)
VG _k	= vehicle group (A, B, C, D)
CA ₁	= car age (0-3, 4-7, 8+)
(In)	= interaction terms
E ijkl	= is assumed normal with mean zero and variance σ^2 .

The fitting method was OWLS see Appendix 1.

The data set was collected from claims occurring from 1 April 1979 to 31 March 80 is Column (1) and (3) of TABLE 7.5 shows for all combinations of rating factors the exposures and actual claims respectively. TABLE 7.3 shows the analysis of variance table.

TABLE 7.3

CLAIM FREQUENCY		ANALYSIS OF VARIANCE:
FITTED BY OWLS	:	DEPENDENT VARIABLE Log P

FACTOR	(1) SUM OF SQUARES	(2) DEGREES OF FREEDOM	(3) MEAN SQUARES (1) - (2)	(4) MEAN SQUARES RATIO (3) - RESIDUAL
COVER	179.2	1	179.2	167.5
CAR AGE	205.0	2	102.5	95.8
VEH. GP	187.6	3	62.0	58.4
P/H AGE	260.3	4	65.0	60.8
COVER -X- CAR AGE	25.2	2	12.6	11.7
COVER -X- VEH. GP	5.0	3	1.6	1.5
COVER -X- P/H AGE	2.4	4	0.6	0.5
CAR AGE -X- VEH. GP	2.2	6	0.3	0.3
CAR AGE -X- P/H AGE	19.3	8	2.4	2.2
VEH. GP -X- P/H AGE	11.6	12	0.9	0.9
RESIDUAL	79.1	_74	1.1	-
TOTAL	977.5	119		

By inspection of the mean square ratio column (4) the most important factor is cover, and the interaction terms are relatively unimportant. Hence the fitted proportions will be modelled using all the main effects. Also notice that for this example 85% of the variation bas been explained by the main effects.

TABLE	7.4	shows	the	estimates	of	parameters.
-------	-----	-------	-----	-----------	----	-------------

TABLE 7.4						
CLAIM FREQUE	NCY	PARAMETER	ESTIMATES			
ADDITIVE	MODEL (M.	AIN EFFECTS	5 ONLY)			
	overall	=	-2.044			
cover	comp	5	0.111			
non	comp	=	-0.244			
car age	0 - 3	=	0.296			
	4 - 7	=	0.044			
	8+	=	-0.209			
vehicle group	Α	=	-0.220			
	В	=	-0.076			
	Ċ	=	0.017			
	D	=	0.351			
policyholder age	17-20	Ŧ	0.687			
	21-24	=	0.487			
	25-29	=	0.180			
	30-34	=	0.042			
	35+	=	- 0.100			

To obtain an estimate of the proportion of claims for a particular cell ijk, with only main effects

r

$$\log \frac{\hat{P}}{1-\hat{P}} = \hat{T} + \hat{CO} + \hat{PH} + \hat{VG} + \hat{CA} \text{ where } \hat{I} \text{ is at least squares estimate}$$
$$= \hat{X} \text{ (say)}$$
$$\hat{P} = \frac{\hat{e}}{1+\hat{e}} \hat{X}$$

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For example from TABLE 7.4 we have for cell definition non comp, car age 8+, vehicle group D, and policyholder age 35+, the following calculations to estimate P:-

overall value	- 2.044
non comp	- 0.244
car age 8+	- 0.209
vehicle group D	0.351
policyholder age 35+	- 0.100
total	2.246

Finally TABLE 7.5 shows the actual verses fitted analysis where:-

col (1) = Exposed to risk col (2) = Fitted proportion of claims (main effects) col (3) = Actual claims col (4) = Expected number of claims col (5) = Actual/Expected col (6) = (Actual - Expected)²/Expected = chi squared statistic

and

....

a agriculta dago	CLAIM FREC	DUENCY DEPE	NDENT VARI	TABL ABLE LOG	E 7.5	s, ma	ACT IN EFFE	UAL V.	FITTED ANALYSIS
COVER	VEH. AGE	VEH. GP	P/H AGE	EXPOS (EX) (1)	FITTED PROP'Ns (2)	CLA ACTL (A) (3)	IMS FITTED VALUES (4) (1)x(2	A/E *100 (5))	$\frac{(A-E)^2}{E}$ (6)
COMP	0-3	A	17-20 21-24 25-29 • 30-34 35+	36 68 123 179 1609	. 230 . 193 . 155 . 138 . 124	5 18 25 29 230	8 13 19 25 199	60 133 131 117 116	1.29 1.51 1.84 .74 4.88
		B	17-20 21-24 25-29 30-34 35+	49 146 349 555 5772	259 225 177 153 140	13 36 63 93 825	13 33 62 88 810	102 110 102 104 102	.01 .30 .03 .34 .29
		С	17-20 21-24 25-29 30-34 35+	49 176 433 863 8165	.278 .242 .191 .171 .152	7 44 79 159 1262	14 43 83 147 1240	51 103 94 107 102	3.22 .04 .14 .77 .40
		Q	17-20 21-24 25-29 30-34 35+	18 119 289 576 4203	.347 .309 .248 .223 .200	7 24 69 135 818	5 37 72 129 840	112 65 96 105 97	.07 4.41 .10 .32 .58
	4-7	A	17-20 21-24 25-29 30-34 35+	47 104 174 274 2545	. 188 . 142 . 125 . 111 . 099	10 23 15 25 243	7 17 22 30 252	113 137 69 82 97	. 15 2. 24 2. 08 . 97 . 29
		B	17-20 21-24 25-29 30-34 35+	135 281 526 1005 10117	. 216 . 185 . 143 . 127 . 113	30 61 82 126 1153	29 52 75 128 1139	103 117 109 99 101	.02 1.58 .59 .03 .17
		c 、	17-20 21-24 25-29 30-34 35+	55 211 505 950 7825	.230 .199 .155 .138 .122	19 39 72 115 1034	15 42 78 131 956	125 93 92 88 108	. 95 . 21 . 50 1. 96 6. 35
		D	17-20 21-24 25-29 30-34 35+	20 103 241 476 2963	. 289 . 257 . 203 . 182 . 163	9 28 49 95 515	6 28 49 87 483	156 101 100 109 107	-1.73 .00 .00 .77 2.17
•	÷+	A	17-20 21-24	18 40	.139 .124	1 3	3 5	40 61	. 90 . 77
			25-29 30-34 35+	94 192 3105	.097 .087 .073	11 17 240	9 17 244	120 101 93	. 38 . 60 . 65
		B	17-20 21-24 25-29 30~34 35+	35 81 241 525 6607	. 169 . 146 . 114 . 101 . 090	6 8 24 57 534	6 12 27 53 592	102 43 87 107 90	. 00 1. 25 . 43 . 28 5. 73
		. C	17-20 21-24 25-29 30~34 35+	20 40 130 284 2800	. 177 . 154 . 122 . 107 . 097	4 6 17 34 262	4 16 31 273	113 97 107 109 96	. 05 . 01 . 09 . 27 . 49
•	ه معدم من الم	D	17-20 21-24 25-29 30-34 35+	2 11 51 111 722	. 124 . 187 . 160 . 145 . 131	0 2 8 21 90	0 2 8 16 94	0 97 98 130 95	. 25 . 00 . 00 1. 50 . 20

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	CIATH EVEN	M: DOM		TABL	E 7.5		ACTI	AL V.	F177F0	ANALYSIS
	orvin surd	DEPE	NDENT VART	ABLE LOG	P,OW	LS, MA	IN EFFE	CTS	FILLU	AIREIGIG
COVER	VEH. AGE	VEH. GP	P7H AGE	EXPOS (EX) (1)	FITTED PROP'N (2)	CLA) ACTL 8 (A) (3)	IMO FITTED VALUES (4) (1)x(2)	A/E +100 (5)		(<u>()</u> (6)
N/COMP	0-3	A	17-20 21-24 25-27 30-34 35+	3 9 6 11 67	. 139 . 113 . 052 . 066 . 084	0 1 0 5	1 1 0 1 6	0 77 0 0 87		1.11 .00 .31 .73 .07
		В	17-20 21-24 25-29 30-34 35+	10 29 26 25 192	.171 .159 .117 .101 .101	1 4 6 4 17	2 5 3 3 19	58 86 196 158 83		. 30 . 03 2. 84 . 35 . 28
		С	17-20 21-24 25-29 30-34 35+	12 30 35 47 267	. 193 . 174 . 132 . 119 . 110	1 5 5 27	2 5 5 29	43 115 108 90 92		. 75 . 12 . 03 . 06 . 20
		D	17-20 21-24 25-29 30-34 35+	6 21 , 15 21 93	. 241 . 227 . 169 . 152 . 145	2 5 4 1 15	1 5 3 3 14	133 105 159 31 111		. 21 . 01 . 88 1. 51 . 16
••	4-7	A	17-20 21-24 25-29 30-34	78 90 97 113 573	.140 .113 .088 .078	11 15 6 11	11 11 9 9	100 142 70 119	_	.00 1.34 .77 .35
	- · ·	В	17-20 21-24 25-29 30-34 35+	- 184 252 315 332 2207	. 162 . 137 . 104 . 092 . 082	-3 25 41 35 37 167	- 30 34 33 35 180	84 119 107 105 93		1.22 .77 1.25 .14 .10 .94
		с 	17-20 21-24 25-29 30-34 35+	114 204 271 306 1585	. 174 . 143 . 113 . 100 . 089	24 27 27 22 128	20 30 31 31 141	121 90 88 72 91		. 88 . 33 . 43 2. 40 . 1. 13
		D .	17-20 21-24 . 25-29 30-34 35+	-45 123 125 144 599	. 225 . 195 . 150 . 133 . 119	12 20 16 15 59	10 24 19 . 19 . 71	119 \$4 \$5 79 83		.35 .64 .41 .92 2.18
	8+	A	17-20 21-24 25-29 30-34 35+	386 365 488 633 37 64	. 115 . 096 . 073 . 064 . 056	49 40 39 35 223	44 35 35 40 212	108 114 110 87 105	. •	. 28 . 69 . 37 . 73 . 56
		- <u>-</u> B	17-20 21-24 25-29 30-34 , 35+	512 630 843 1100 6470	. 131 . 110 . 033 . 073 . 065	77 71 81 84 507	67 69 70 81 418	115 102 115 104 121		1.44 .04 1.64 .13 19.00
		C	17-20 21-24 25-29 30-34 35+	179 272 357 478 2608	. 141 . 119 . 090 . 080 . 070	29 40 33 41 205	25 32 32 38 183	115 124 103 108 112		.57 1.85 .02 .24 2.54
		D	17-20 21-24 25-29 30-34 35+	45 77 102 145 549	. 182 . 155 . 119 . 106 . 095	9 13 14 21 55	8 12 12 15 52	110 109 115 137 105		. 09 . 09 . 28 2. 06 . 16
TOFAL	•	•		1030578		11548	11310			

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TEST STATISTICS ON 102 DEGREES OF FREEDOM' (ASSUMING MAIN EFFECTS ONLY USED) TOTAL COL (9) UNWEIGHTED CHI-SQUARE # 117,11 The Actual/Expected statistic is given since it is the standard actuarial test, but it is difficult to interprete for this data. It is found that the chi-squared statistic gives a better overall measure. Only one cell stands out as being very perculiar namely, non comprehensive, vehicle age 8+, vehicle group B, and policyholder 35+. The chi squared statistic = 117.11 with 119 degrees of freedom, this shows a good fit (compare with the Expected Values of a chi-squared statistic E $(\chi^2_{119}) = 119$).

The main effects estimated will be inputted as the frequency data for the premium calculation (TABLE 7.5 column (2)).

7.3.3 Future Work

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The main problem which has not been addressed, is, how well does the model predict numbers of claims (a) overall (b) within portfolio?

As commented in Section 7.2, it is felt that overall levels cannot be statistically predicted.

The within portfolio problem was considered by Coutts(1975) and the additive model with main effects was found to predict reasonably well. However, no other published work has been found. Three methods suggest themselves to judge forecasting methods:-

- (i) compare actual claims, with predicted claims, using past data to estimate parameters, and the present exposures;
- (ii) compare parameter estimates of rating factors over time;
- (iii) formally include in the model a time factor using some form of time series model.

7.4 MATERIAL DAMAGE AND MISCELLANEOUS: Overall

For Accidental Damage (or Fire and Theft for Non-Comprehensive), Third Party Property Damage and Miscellaneous Costs, the object is to obtain the average cost at the mid point of the third quarter 1980 if settled immediately. This value is then projected forward after taking into account the date at which the claim is expected to be notified and the average settlement date. In practice, inflation is assumed to be the main factor for projecting overall levels. This is supported by Ziai (1979). He fitted various probability distributions to several comprehensive AD claims cost data sets. Various techniques were used to estimate the parameters for each of the probability distributions investigated. After estimating the parameters, claim cost distributions were generated by simulation techniques and goodness of fit tests were applied. Next Ziai, investigated the prediction ability of these distributions. The parameters of the original data were adjusted for one year's inflation and, conditional on the known numbers of claims one year later, a simulated predicted claim distribution was obtained. This was compared with the actual claims cost.

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For distributions log normal, gamma and the inverse Gaussian, reasonable predicted results were obtained.

Ziai's work was performed on a relatively small data sets. Since the same Insurance Company supplied both Ziai's data and the data analysed in this thesis, Ziai's results are considered representative of the data used in this thesis. Further even though Ziai's work only considered comprehensive AD, it seems reasonable to use inflation as the basis for projection for all material damage costs. The actual rates of inflation are discussed in section 9.2. Therefore it is assumed that the overall levels of material damage data costs (i.e. AD and TPPD) will reflect the actual average values for the period 1 April 1979 to the 30 March 1980, adjusted for inflation to bring them up to the start of the projection period i.e. 15 August 1980. As far as miscellaneous costs, nominal figures of £2.50 for comprehensive and £1.50 for non comprehensive were used.

Finally, some analyses relating to average settlement have been performed and the following average dates to settlement are being used. For both comprehensive and non-comprehensive AD. 3 months, TPPD, 6 months and M 3 months.

7.5 MATERIAL DAMAGE: Within Portfolio

7.5.1 Review

Very little work has been published analysing claims cost data sets. During the 1960's when claim frequency models were being developed and discussed, no published papers were found discussing in an in-depth analysis of claims cost. Bailey and Simon (1960) and Seal (1968) implicitly discussed claims cost since they analysed claim ratios. Why claims cost were ignored is not clear and, some 15 years on, it is only possible to hazard a guess. A plausible reason is that, as the 1960's inflation was less than 6%, it seemed likely that all the variation in claims experience was attributed to claim frequency. Another factor as far as the UK was concerned, was that until the end of the 1960's, there was a motor insurance tariff in operation. This was organised by the Accident Office Association (AOA) which collected statistics by one way rating factors, so that no tariff office needed to analyse their own data in depth. When the AOA tariff collapsed and the British Insurance Association (BIA) took over the data collecting through the Motor Research Statistical Bureau (MRSB) in the earlier 1970's, companies which contributed to the MRSB were forced to set up data bases which allowed an in-depth claims cost analysis to be studied. During the 1970's, inflation became a very significant factor and though the MRSB analysed their results were confidential and in-depth claims cost; were not published. Johnson and Hey (1971) did not analyse claims cost explicitly, however, their model is being used by the MRSB to smooth claims cost data. Kahane and Levy (1975) analysed Israeli claims cost data, using an additive model but their analysis is only briefly explained.

Baxter and Coutts (1977) and Baxter Coutts and Ross (1980) analysed the same set of comprehensive AD data using the model suggested by Johnson-Hey (1971) and the results were compared with OWLS. The dependent variable was the average AD cost per intimated

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claim, the independent variables were the rating factors, of policyholder age, car age and vehicle group and the weights were the number of intimated claims.

The results using both models were very similar. In addition Baxter and Coutts (1977) quoted the analysis of variance table, which showed that the main effects explained about 70% of the variation and that the interaction terms could be ignored for fitting purposes. Baxter and Coutts (1977) started to consider the residual analysis and Baxter Coutts and Ross (1980) continued this analysis and confirmed the earlier results and the validity of the underlying assumptions were satisfied viz:

the additive model with main effects gave a reasonable fit
the variance was proportional to the inverse of the number of intimated claims.

However, two major errors were made by Baxter Coutts and Ross (1980), and so both these conclusions have to be considered suspect. The errors were:-

(a) If the model and variance assumptions were correct then a residual analysis showing fitted values against standardised residuals would give a plot of points which would look random Anscombe and Tukey (1963). It was brought to the authors' notice that there was a trend in the residual plot, as the fitted values became larger, so did the standardised residuals. This implied that the underlying

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assumption of the variance being proportion to the inverse of the number of intimated claims was suspect. This was discussed with Nelder (1980) who suggested using a gamma error distribution. However, this also failed to give a satisfactory result.

(b) The GLIM package procduces an estimate of the residual means square, which has expected value 5^{2} (the error variance) for an adequate model. In addition any statement concerning the percentage of variation explained by the analysis of variance depends on the estimate of σ^{z} being reasonable. For the main effects model suggested by Baxter and Coutts (1977), the residual mean square was equal to 82,000 units. However, an independent estimate of σ^2 obtained from the data of individual claims, estimated σ^{-1} to be 43,000 units, approximately half that assumed by the model. Hence it appears that the main effects model is not entirely adequate. It was hoped that by taking logs of the individual claims (i.e. assume a log normal distribution) this might overcome this problem, but this proved unsatisfactory.

The ASTIN Netherland Group (1982) also analysed material damage claims cost and used an additive model, similar to the Johnson-Hey (1971) model. They did not test the underlying assumptions and in particular an independent estimate of residual sum of squares was not calculated. So it is felt that given the preceeding comments these results must be suspect.

7.5.2 The Analysis

The data collected for AD, TPPD was typically unbalanced, with a number of cells giving zero claims cost. Four separate analyses were performed on comprehensive, non-comprehensive AD and TPPD. For comprehensive AD the actual average costs and number of claims are listed in TABLE 7.6 columns (1) and (4) respectively.

To smooth the data OWLS was selected, even though it is certainly not a good model. It was considered better to use a smoothing technique rather than either the actual averages or one way tables. In TABLE 7.6 column (2) shows the comprehensive AD values, only main effects. Column (3) shows the difference between actual and smoothed values. Whilst column (5) shows the count times column (3). By inspecting column (5), it is obvious that the fit is not especially good but, overall the sum of column (5) = \pounds 3504. Details are only shown for comprehensive AD, the other material costs being smoothed in the same way.

The smoothed results inputted are shown in the premium calculation input TABLE 10.1. For convenience TPPD value will include the M cost.

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		A D	TA COMPRESENSING SMOOTH	BLE 7.6		:	•
		ΛV	COMPREMENSIVE SMOOTH	ED VALUES	ACTUAL V. FITTED	•	
VEH AGE	VEH GROU	P/H AGE	(1) ACTUAL	(2) Expected	(3) ACT-EXP	(4) Count	(5) CNEXLIEF
0-3	٨	17-20	455.37	416.70	70 47		
		21-24	320.65	302.75	17.70	19.00	173.34
		25-29 30-34	255.86	302.13	-46.27	20.00	-1225.52
		35+	202.36	257.06	25.17 15.30	20.00 239.00	204.63 3657.70
	В	17-20	423.69	421.11	2.50	.15.00	30.77
		21-24	312.26	307.36	4.90	37.00	191.21
		30-34	204.00	306.54 262 96	1.34	60.00	91.40
		35+	269.18	271.46	-2.20	352.00	-1958.00
	С	17-20	472.01	468.10	3.91	2.00	35.10
		25-29	212.04 366.93	354.35	-31.01	10.00	
		30-34	206.72	309.25	-22.53	170.00	1112.20
		35+	325.57	318.45	7.14	1320.00	0412.46
	D	17-20	1638.30	563.71	1067.59	9.00	9626.35
		25-29	463.63	454.76	202.02 9.49	27.00 78.00	0204.03 746 SC
		3034	468.53	409.00	58.67	147.00	0424.53
		35+	394.99	419.06	-24.07	872.00	-20768.09
4-7	Α	17-20	225.98	373.42	-147.44	12.00	-1769.22
		21-29	1/0.20	257.67 252 94	-34.47	23.00	-1942.72 -AFO F/
		30-34	147.72	214.57	-12.05	27.00	-1204.93
		35+	237.56	223.77	13.79	266.00	3465.37
	B	17-20 21-24	345.05 and 10	377.82	-32.77	39.00 27.00	-1245.20
	-	25-29	247.06	263.25	-16.12	2,.0∖ 72.00	-1407 .45
		30-34 35+	244.14 231.62	210.97 228 .17	23.17 3.45	133.00 1254.00	3007.41 4321.10
	с	17-20	250.16	424.81	-174.65	21.00	-3667,73
		21-24	346.63	311.00	35.57	41. 00	1430.10
		25-29 Xo-74	284.90	310.24	-25.34	35.00 100 oo	-2154.19 ZOKO BA
		35+	270,77	275.17	-4.40	1123.00	-4933.49
	D	17-20	272.59	525.42	-252.83	2.00	-2275.47
		21-24	330.07	411.67	-01.60	31.00 57 60	-2020.01 5544 44
		20-27 30-34	298.31	366.57	-68.26	104.00	-7029.40
		35+	374.87	375.77	90	586.00	-528.47
ú+	А	17-20	308.37	307.20	1.17	2.00	2.33 -118-05
		21724	191 68	172.63	- 17.00	14.00	-13.36
		30-34	69.06	148.36	-79.30	20.00	1505.97
		35+	132.80	157.56	-24.76	263.00	-6511.55
	B	17-20	118.74	311.61	-192.87	9.00	-1735.83
		21-24 25-29	274.81	197.04	-79,93	25.00	-1979.48
		30-34	172.46	152.76	19.70	65.00	1299.96
		35+	181.25	161.96	19.29	619.00	11938.25
	C	17-20	241.51	358.60	-117.09	4.00	-468.37
		21-24 25-29	207.09	244.03	127.91	20.00	2558.14
		30-34	246.63	199.76	46.37	30.00	1701.16
	•	35+	186.35	208.96	-22.61	304.00	-6872.63
	D	17-20	107.21	459.21 345 AK	-352.00	1.00	- 352.00
		25-29	247.41	344.64	-97.23	9.00	-875.06
		30-34	414.32	300.36	114.46	27.00	3090.34
		35+	312.60	307.56	5.04	97.00	340.65
		· , ·			TOTAL		-3504.06

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-3504.06

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7.5.3 Future Work

The analysis of claims cost is still relatively new. In particular some work establishing the error distributions by cell is necessary so that reasonable models can be used. The GLIM package should be of great assistance for the researcher. Possible lines of research could be to work on using the gamma or inverse Gaussian error distribution.

The analyses so far have explicitly ignored the variable time, which can be interpreted as explaining inflation. Dawid (1981) suggested the following generalised model:-

•	let log AD	Ξ	Nt
		+	Q _i +R _j +S _k
		+	interaction terms involving time
		+	interaction terms not involving time
		+	error not including time
	,	+	error including time
	where AD is	the av	verage AD cost

 N_t is the main effect of time and Q_i, R_j, S_k are the underwriting factor effects

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This model is known as a mixed model, where the mixture is between the fixed effects rating factors Q,R,S, (which are not random) and the random effect time N. Also there are two variances to estimate. The statistical problem lies in estimating these variances. Standard methods usually produce one negative variance, which is difficult to explain (Harville 1978).

Interpretation of the estimate is:

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(a) N_t is the effect of inflation and can be compared with the inflation used in past;

(b) the interaction terms including time relate underwriting rating factors to time after eliminating inflation. If these interactions are important then it will be difficult to predict costs using this model.

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CHAPTER EIGHT

THIRD PARTY BODILY INJURY (TPBI) COSTS

8.1 Introduction

The hardest part of the statistical analysis, since, on average, only about 5% to 7% of all claims per year involve bodily injury costs, yielding for the example portfolio less than 1,000 TPBI claims but this represents about 20% of the total cost. However, before discussing details, it is necessary to look at the arguments against pooling data from different companies to arrive at an input to premium calculations.

Since the number of claims for each company is, in practice, small, it seems reasonable that all companies should pool their data. That would not, however, solve the problem, for a reason best illustrated by way of an example. TABLE 8.1 shows some typical average bodily injury costs per claim for an individual company. They have been adjusted for earnings inflation plus 'judgement drift' (Section 3.1) to bring them all up to 15 August 1980, the projected date.

TABLE 8.1

BODILY INJURY COSTS BY YEAR

Year of	Average Bodily Injury Costs per Claim
Accident	Inflated up to 15 August 1980
	£
1972	73
1973	50
1974	70
1975	35
1976	65
1977	52

If inflation were the only factor operating on these averages, relatively constant values would be expected assuming that inflation and 'judgement drift' assumptions were correct. However, it is clear that there is wide variation. This is due to the extreme skewness of the underlying distribution of bodily injury claims; there is no reason to suppose that the distribution shifts from year to year.

Corresponding pooled data for all companies are available in the UK from the MRSB but are not available for publication. However, if they were, then the sample size would be far greater and the variability smaller than for any single company. Suppose an average claim of £58 (in 1980 values) resulted from pooled data for years 72 to 77. If the Company based rates on this, a large positive surplus would be shown if the 1975 experience were repeated, or a large negative surplus if the 1972 figures occurred again. The real problem is that the Company experiences a small sample of the total market experience and therefore its premium rates should make an allowance for its own variability. How this is to be accomplished in the model is not obvious. Theoretically, a factor for the variance can be included, (Kahane and Levy (1975)), but the rates obtained may be too uncompetitive.

Alternatively, it could be included explicitly in the claim reserving philosophy. Additional reserves could be set up in 'good' years which will be released to supplement the claim experience in bad years. This reserve is often referred to as the equalisation or catastrophy reserve e.g. Trayhorn (1980). In Finland this reserve

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is allowed to be set up, and can offset positive surplus, so reducing the tax liability. However, in the UK it has to be met out of profits after tax, and so is not popular.

8.2 Overall

The Group has to decide the basis of projecting bodily injury claims. Assuming that they have decided to use their company's data, they have to select the base period. The possibilities are:-

- (i) using the latest results (1978 and 1979), which, however,entail predominantly manual estimates;
- (ii) combining claims experience from earlier years, which will reflect 'good' and 'bad' years experience. The results will contain some manual estimates but these are expected to be realistic and more reliable than those estimates contained in method (i); as older estimates will be based on claims 3 or more years old and the claim assessor is expected to establish reliable liability costs.

There is no correct answer to the problem of prediction and the Group has to arrive at a view, and hold it until evidence is produced to indicate that the view has to be altered. A choice between methods (i) and (ii) hasto be made bearing in mind that:

- method (i) contains the latest information, but has inherent subjective estimates.
- method (ii) is more objective than method (i), but it uses data which are 3 or more years old to predict claims cost 4 years into the future.

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Due to the uncertainty about the most recent claim estimation procedures, method (ii) was accepted in preference to the more subjective approach.

The following data were collected by year of accident to be used as the base period. It was also decided on statistical evidence that comprehensive and non-comprehensive data were to be shown separately.

- (A) The average TPBI claims cost per intimated (payments plus outstanding) for years of accident 1972 to 1977.
- (B) All averages calculated in (A) are discounted to allow for the time to settlement, i.e. to estimate its value if the claim was settled immediately. Then this value is inflated to the start of the projection period, i.e. 15 August 1980.

The method of adjustment is best explained by an example. For the average TPBI cost for 1972, it is assumed it takes two years to average settlement (Section 3.1), therefore, the average TPBI is to be discounted to 1st July 1972. This value is then inflated from 1st July 1972 to the mid point of 3rd quarter 1980. Since it is assumed that the rate of discount for settlement is the same as the inflated value, it is only necessary to inflate 1972 from 1st July 1974 to 15th August 1980 and 1973 from 1st July 1975, etc. The inflation rates used are detailed in Chapter 9 and are based on the BIA Economic Advisory Group (EAG) quarterly reports.

8.2.1 Comprehensive Analysis

The following table summarises the data collected for comprehensive:-

TABLE 8.2

MOTOR COMPREHENSIVE

THIRD PARTY BODILY INJURY CLAIMS COST ANALYSIS

	(1) Tafiation	(2)	(3)
	Adjustment	Actual Average	
Year of	to	TPBI per	Inflation adjusted
Accident	15 Aug. 80	intimated claim	TPBI col. $(1)x(2)$
<u></u>		£	£
72	2.587	28.1	72.7
73	2.127	23.3	49.6
74	1.755	40.0	70.2
, 75	1.548	22.4	34.7
76	1.377	47.2	65.0
77	1.197	43.5	52.1
72-76			58.1
72-77			57.2

As has been pointed out above, the variation by year is due in part to:-

- (a) the underlying skew third party bodily injury distribution;
- (b) the small database;
- (c) the use of past inflation rates based on Economic AdvisoryGroup data which may not be appropriate, in particular theeffect of 'judgement drift' is very difficult to estimate;

- (d) the proportion of TPBI injury claims is very small, about
 5-7%, and a small deviation in this proportion will affect
 the intimated average;
- (e) The effect of large claims on the average (see section 8.3.2).

These problems make selection of the base period difficult. Certainly the use of one year is not advisable. If this is rejected the decision of what years to include has to be made.

From TABLE 8.2, it seems that the 'good' years are 1973, 1975, and 1977, whilst the bad years are 1972, 1974 and 1976. Combining 'good' and 'bad' years seems a reasonable approach (as with constructing a standard life table). It is suggested that two combinations of years are selected, 1972 to 1976 and 1972 to 1977. The values were calculated by weighting the inflated averages by the number of intimated claims. The respective values being £58.1 and £57.2. For comparison purposes, the actual average TPBI claim cost per intimated claim, based mainly on manual estimates and unadjusted for inflation for 1978 and 1979 at 1st July 1980, were both £47.00.

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TABLE 8.3 summarises the data for non-comprehensive.

TABLE 8.3

MOTOR NON-COMPREHENSIVE

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THIRD PARTY BODILY INJURY CLAIMS COST ANALYSIS

	(1)	(2)	(3)
	Inflation		
	Adjustment	Actual Average	
Year of	to	TPBI per	Inflation adjusted
Accident	15 Aug. 80	intimated claim	TPBI col. $(1)x(2)$
		£	£
72 ·	2.587	76.4	197.6
73	. 2.127	72.6	154.4
74	1.755	86.2	151.3
75	1.548	110.0	170.3
76	1.377	75.4	103.8
77	1.197	82.0	98.2
72-76			155.0
72-77		• .	145.7

Similar comments to those made to comprehensive can be made about the variations in averages (Section 8.2.1)

To show consistency between comprehensive and non-comprehensive two combination of years are selected, 1972 to 1976 and 1972 to 1977 the respective values being £155.0 and £145.7. For comparison purposes, the actual average TPBI claims cost per intimated claim, based mainly on manual estimates, unadjusted for inflation for 1978 and 1979 at 1st July 1980 were £163 and £136 respectively.

Finally, after some data investigations the evidence suggests that on average the assumption of a 2 year period to settle was reasonable for both comprehensive and non comprehensive.

8.3 Within Portfolio Analysis

8.3.1 Review

The small data-base, when sub-divided even further will make any formal statistical analysis difficult. This is not to say it should not be undertaken. Little (1978) has used generalised linear models to analyse small data-bases, where the underlying distributions are not necessarily normally distributed. Papers have been published in the past few years giving details of some analysis of TPBI claims outside the UK, but all are based on large data bases. Chang and Fairley (1980) investigated the relationship of TPBI costs from the State of Massachusetts using the rating factors, type of driver and district within State. The emphasis of this paper was on the difference in fit between the additive and multiplicative models. The ASTIN Netherlands Group (1982), performed an analysis of third party data by using weighted least squares. The rating factors analysed were weight of vehicle and car type. They came to the conclusion that weight of vehicle was more important.

A general criticism of Chang and Fairly (1980) and ASTIN Netherlands Group (1982) were that no investigation of the underlying distribution was undertaken and no residual tests comparing fitted against actual were published to support the assumptions of the models. In addition no independent estimate of the residual sum of squares from the data was calculated, Section 7.5.1, to judge whether the inferences made about the analysis of variance tables were reasonable.

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Hallin and Ingelbleek (1981) work was non-parametric. They analysed third party costs collected by all the Swedish Insurance Companies. The emphasis of the work was to establish an ordering of several rating factors. Westenberger (1983) on the Swedish data has shown a contrary result, namely that the third party cost are approximately proportionate to the exposure. This data were also analysed using GLIM and Shrewsbury (1983), supported Westenberer's results. There was evidence to suggest that the rating factors can be ordered but they do not explain a statistically significant part of the variations, hence inference concerning the ordering is not all that important.

8.3.2 Large Claims

As the data are broken down into even smaller groups, the effect of large claims becomes very important, as a large claim in a small cell will disproportionately affect the results. Hence it seems reasonable to apply some method of smoothing has to be applied. In the example, all claims over £10,000 were cut off at that value and the excess was respread over the whole portfolio.

To attempt to judge the effect of this crude truncation on the overall yearly results, TABLE 8.4 has been prepared for comprehensive.

TABLE 8.4

MOTOR COMPREHENSIVE

THIRD PARTY BODILY INJURY CLAIM COST ANALYSIS

LARGE CLAIM ADJUSTEMENT

	(1) Number of	(2)	(3)
Year	large claims in excess £10,000	Average TH Intimated clai	PBI per im inflated
	<u> </u>	large claims	large claims
		Unadjusted	Adjusted
		£	£
72	2	72.7	70.0
73	6	49.2	43.8
74	10	70.2	41.2
75	7	34.7	30.0
76	12	65.0	42.0
77	10	52.1	35.4

The effect of removing the claims in excess of £10,000 reduces the variation between years, except for 1972. A more reasonable method is to decide on the excess point for 1972 and inflate all subsequent years by the assumed inflation rate. However, claims below £10,000 were not available hence this method could not be applied. Another approach would be by taking into account inflation and the underlying distribution e.g. Ziai (1979). The truncation method is applied by the MRSB on its claim analysis as it is easy to apply.

These two major statistical problems alone seem to lead one to reject any in-depth analysis on small data. If this approach is followed, then an overall TPBI cost for comprehensive and non-comprehensive should be used in the office premium calculation and the resulting premium structure would reflect claim frequency and material costs.

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This defeatist approach was rejected and some attempts to analyse one-way rating factors were undertaken. If it had failed then it was always possible to revert to an overall value but the empirical approach did seem to give reasonably accepted results.

All discussion concerning large claims assume no reinsurance. For the Insurance Company the largest claim made was £100,000 which is well below its retention level. Hence the cost of reinsurance was ignored.

8.3.3 A Practical Approach

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The following was the rationale for the proposed analysis.

Even though the overall levels of average TPBI costs vary by year, it would seem reasonable that when comparing several years the relative cost by policyholder age to the overall value should not fluctuate greatly between years. In addition to help reduce the variation of the relativities, claims in excess of £10,000 were truncated at £10,000.

The relative value obtained would then be multiplied by the overall input value, as detailed in Section 8.2. The following example in TABLE 8.5 for comprehensive, policyholder age 35+ give the details of the calculation.

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TABLE 8.5

MOTOR COMPREHENSIVE : TPBI WITHIN PORTFOLIO EXAMPLE

POLICYHOLDERS	AGE	35+ -	RELATIVITIES
		521	

Year of Accident	(1) Claims Cost after large claim adjustment No inflation	(2) Overall Average after large claim adjustment No inflation	(3) (1) - (2) Relativity
72	26.2	27.0	•97
73	20.7	20.6	. 1.01
74	22.4	23.5	.96
75	21.1	19.4	1.09
76	29.5	30.5	•97
77	31.4	29.6	1.06
72-76			1.00
72-77			1.01

The view was taken that the variation between years was reasonable. The input value for 35+ for 72-76 would be 1.00 x 58.1 = 58.1 and for 72-77 would be 1.01 x 57.2 = 57.8 (from TABLE 8.2).

TABLE 8.6 summarises all the relativities:-

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TABLE 8.6

MOTOR COMPREHENSIVE TPBI WITHIN PROTFOLIO : SUMMARY & RELATIVITIES

Year of <u>Accident</u>	17-20	Policy 21-24	/holder / 25-29	Age <u>30-35</u>	<u>35+</u>
72	1.29	3.06	1.38	0.60	0.97
73	.24	3.31	0.84	0.76	1.01
74	1.55	1.99	1.18	0.94	0.96
75	1.13	.67	0.64	1.60	1.09
76	.18	.45	2.04	0.82	0.97
77	5.59	.98	1.38	0.63	1.06
72 - 76	.92	2.14	1.17	0.94	1.00
72 - 77	1.60	2.03	1.19	0.90	1.01

The small number of claims generate a wide variation of relativity by year in the younger ages. Apart from 17-20 which is based on less than 200 intimated claims per year the trend is that the relative costs for TPBI reduces as the policyholder gets older. TABLE 8.7 gives the input values i.e. relativities times values in TABLE 8.2.

TABLE 8.7

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INPUT VALUES : TPBI

COMPREHENSIVE

Paga	Age	17-20	21-24	25-29	30-35	35+
72 - 76		53.63	124.45	67.81	54.90	58.14
72-77		91.50	116.16	68.03	51.63	57.81

TABLE 8.8 shows the relativities for non-comprehensive.

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TABLE 8.8

MOTOR NON-COMPREHENSIVE TPBI WITHIN PORTFOLIO : SUMMARY

	RELATIV	ITIES			
Year of		Policy	yholder <i>l</i>	Age	
Accident	17-20	21-24	25-29	<u>30-35</u>	<u>35+</u>
72 73 74 75 76 77	1.37 1.28 1.71 1.38 2.27 1.96	1.04 0.80 1.19 1.33 1.58 1.17	0.88 0.84 1.33 1.16 0.64	0.52 1.02 1.01 0.32 0.57 0.22	1.06 1.10 0.69 0.96 0.84 1.04
72-76 72-77	1.59 1.64	1.17	0.97 0.92	0.69 0.61	0.92 0.94

The results are not as well behaved as for comprehensive for example, the relativities reduce for age 30-34. This might be expected as the number of claims are small. TABLE 8.9 gives the input values i.e. relativities times values in TABLE 8.3.

TABLE 8.9

INPUT VALUE TPBI

NON COMPREHENSIVE

Rase	Age	17-20	21-24	25-29	30-35	35+
72 - 76		245.26	181.13	150.66	106.79	142.58
72-77		238.97	170.43	133.93	89.44	137.35

This completes the statistical analysis of TPBI costs.

TABLE 8.10 summarises the analysis performed for the claims proportions and claims cost

TABLE 8.10 - SUMMARY OF ALL CLAIM EXPERIENCE ANALYSES

		Whether Modelling	-
Type of Analysis	Rating Factors	Technique used	Overall Level
Claim Proportion	All	YES	1978
Accidental Damage	All	YES	Inflation
Property Damage	All	YES	Inflation
Miscellaneous	Comp. & Non-Comp.	NO	Inflation
Bodily Injury	Comp., Non-Comp. Policyholder age	YES	1972-76 1972-77) + Inflation)

This ends the claim experience analyses.

CHAPTER NINE

ECONOMIC FACTORS

9.1 Expenses

It is not proposed to go into detail about the collection of data to obtain a breakdown of expenses. A paper by Rushton (1977) has described the process very well. As far as the Group is concerned, it will obtain this information directly from the accounting department within the Company, which projects future expenses. The only factor it would consider is the future rate of inflation related to expenses (see Section 9.2).

As far as the premium formula is concerned, two approaches are available.

The first is to consider all expenses as fixed in the short term - equation (4.2). The total value of expenses is obtained from the accountants and then expressed as a percentage of premium income. Commission and other premium related expenses are added to this percentage. In practice, the value should be between 25% to 45%, depending on type of company. The rationale behind this is that in the short term (say 2 years) staff levels (which make up 80% of expenses) are virtually fixed. Thus, it is argued, further sophistication is needless. The fixed expense level was calculated to be 45% of premium for the example under consideration.

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The second method equation (4.3) is to divide up costs into those which are fixed and expressed as a percentage of premium, and those which are identifiable as separate totals such as claim cost expense, new business expense, lapse expense and endorsement expenses.

For the example, the following were used and then inflated to 1980 levels:-

Fixed as a percentage of premium (incl. commission) 17% Claim Cost £14 per claim New Business £6 per Policy Lapse £1.50 per Policy Renewal £1.50 per Policy £1.50 per Policy £2.25 per Policy

Both methods will give the same overall expense allocation, if the portfolio is similar to that of the base period. However, if the portfolio changes, the allocations will change. The choice between the two methods depends on the accounting methods of the company. Section 12.2 will show that the different allocations effect the premium structure significantly.

9.2 Inflation

The Group will have to form some overall views on inflation. As mentioned earlier, general economic factors and government policies will dominate. Usually, several scenarios will be followed. Different rates of inflation will be applied to different parts of the analysis. In this respect, in the UK, the B.I.A. Economic Advisory Group (EAG) is helpful in

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reporting and projecting inflation separately for each type of claim. Another problem is to establish which index will be an appropriate indicator for each type of inflation risk. For purposes of this example the EAG results were followed.

TABLE 9.1 is a summary of the inflation rates and inflation indices used to project costs:

TABLE 9.1 - INFLATION SUMMARY

Future Inflation 1980 until 1984

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INFLATION INDEX HIGH LOW

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Material Damage Claims	Earnings + Material Goods	11%	7%
Bodily Injury Claims	Earnings + Judgement Drift	17%	13%
Expenses	Internal, based on Salary Increase	10%	10%
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TABLE 9.2 shows the assumed previous inflation rates to adjust past claims up the projection date 15 August 1980.

TABLE 9.2

PAST INFLATION RATES

AD, TPPD, M (BASED ON COMPANY DATA)

Year	Quarter	Inflation per annum
79	2	14%
79	3	14%
79	4	17%
80	1	17%
80	2	15%

TPBI (BASE ON EAG FIGURES)

Year		Inflation per annum
1974		17.3%
1975		26.1%
1976	Ν.	16.5%
1977		10.3%
1978		14.6%
1979		15.5%
1980		18.8%

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Notice in TABLE 9.2, that the past inflation adjustments for material costs are significantly higher than the forecast values. The reason is that the future rates of inflation are expected to fall dramatically in 1980 through to 1981 and beyond.

The following method was applied to the TPBI rates in TABLE 9.2 to obtain the past inflation adjustments used in TABLE 8.2.

Assume that 1972 TPBI claims occurred on 1 July 1972. This value has to be inflated to the start of the projection period i.e. 15 August 1980. As equation (4.1) takes into account the 2 year average date to settlement, the 1972 projected value has to be discounted to allow for this. Therefore the inflation adjustment is from 1st July 1974 to 15th August 1980 that is in terms of inflation:

 $(1.173)^{\frac{1}{2}}$ (1.261) (1.165) (1.103) (1.146) (1.155) (1.88) = 2.587

TABLE 9.3 gives the full set of TPBI adjustments

TABLE 9.3

	TPBI	INFLATION	ADJUSTMENTS	FOR	YEARS
Years		Ad.	justment		
1972	`	2	.587		
1973		2	. 127		
1974		1	• 755		
1975		1	•548		
1976		1	•377		
1977		1	.197		

This ends the economic assumption discussion.

CHAPTER TEN

THE CALCULATION OF PREMIUM RATES

The calculation to arrive at an office premium was performed on a microprocessor by applying equations (4.2) and (4.3) and the input data described in chapters 7, 8 and 9. All the programs used were written by personnel employed by the Insurance Company.

All the claims cost data were adjusted to bring them up to the inflation levels assumed at the start of the projection period i.e. 15 August 1980. TABLE 10.1 shows for all rating factor combinations, the input data namely:

- Col (1) Claim proportion
- Col (2) AD cost
- Col (3) TPBI cost
- Col (4) (TPPD + M) cost

These values then have to be adjusted by inflation according to the rates applicable in TABLE 9.1 to the premium date 1 October 1981 as per section 6.1. This is automatically performed by the computer.

TABLE 10.1 OFFICE PREMIUM CALCULATION

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						35+	.1	310	307.06	33.14	403.21		1 I W . W /

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TABLE 10.1 OFFICE PREMIUM CALCULATION

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COVER	VEH.	AGE	VEH. ORD	P/H ACE	CLA	1M	AL	1461	(TPPD+K)		OFFICE
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NON-COMP	0-3		A	17-20	.1	519	31.35	24. 21	7.7 71.		PREMIUM
				21-24	. 1	130	37.72	101.1.5	ر د در د		115.69
				25-22	.0	520	50.50	1.0.60	14. J/		77.10
				30-34	.0	660	54.155	166.79	' <u> </u>		29.37
				35+	.0	840	44.74	140.50	71 22		27.60
			В	17-20	. 1	710	23.69	245.00	17 142		47.00
				21-24	1	590	30.24	181.13	40 64		139.13
				25-29	.1	170	40.64	150.74	21.01		105.53
				30-34	. 1	010	44.82	106.72	17 94		65,02
				35+	. 1	010	35.08	142.58	71 77		43.42
			С	17-20	. 1	930	36.57	245.26	28.10		04.63
				21-24	.1	740	45.14	101.13	23.12		103.63
				25-29	.1	320	55.52	150.66	17.57		116.97
				30-34	. 1	190	59.77	106.79	3.12		74.45
				35+	.1	100	42.96	142.58	22.00		52.17
			D	17-20	. 2	410	63.10	245.26	52.95		67.44
				21-24	.2	270	71.67	181.13	57.99		224./1
				25-27	- 14	600	32.05	150 74	29 75		177,52
				30-34	. 1	520	86.30	106.70	20.00		
				35+	.1	450	16.49	142.53	46.02		
	4-7		A	17-20	. 1	400	35. 03	245.26	A2 02		70.00
				21-24	.1	130	43.61	1:11 13	67 10		
				25-29	.0	800	53.98	150.66	47 42		37.74
				30-34	.0	7:00	50.23	106 79	47.40		
				35+	.0	710	4:1.42	142.50	55.0°		シア・アイ ハー ロン
			В	17-20	- 10	520	23. 17	245.26	62 53		44.22
				21-24	. 1	370	33.25	451.17	67 60		
		٠.		25-29	- 10	040	44.32	150.76	47 99		17.37 20.37
				30-34	. 9	920	40.57	106.70	47.50		64.07 AL. 77
				35+		320	33.76	142.58	56 46		
			с	17-20	. 1	740	40.25	245.26	52.85		154 06
			·· -	21-24	. 1 -	400	40.02	101.13	57.69		100.85
				25-29	.1	130	52.20	150.66	38.26		70.91
				30-34	.10	000	63.45	106.79	32.05		50-17
				35+	.0	370	53.64	142.50	46.73		54.53
			D	17-20	.2	250	66.73	245.26	77.63		224.03
			-	21-24	. 1/	950	75.35	131.13	82.72		164-92
				25-29	. 15	500	05.73	150.66	63.03		110.66
				30-34	. 1	330	89.90	106.79	57.60		81.32
				35+	.1	190	30.17	142.58	71.55		36.03
	<u>.</u>			17-20	- 1	150	15.36	245.26	63.32		90.60
	8+		A	21-24		26.0	23.94	101.13	63.36		67.93
			•	21 27	0	77.0	34.31	150.65	40.72		_ 43.81
				23-27	.0	540 Í	33.57	106.79	43.32		30.36
				30-34	0	560	28.76	142.50	57.19		32.78
			-	17-20	•••	310	5.70	245.26	63.32		109.97
			В	17-20	1	100	14.20	181.13	59,86		75.85
				21-24	 	230	24.65	150.66	49.23		43.27
				20-29		730	20.91	105.79	43.83		33.27
				30-34		450	1.2.10	142.53	57.70		36.04
				30+		410	20.58	245.26	54.09		119.56
			С	17-20	· · · · ·	1420	20100	101.13	59.13		83.07
				21-24	<u>ل</u> ه.	170	22.10	150.60	39.50		53.11
				25-27	••	200	Z	105.79	34.09		37.14
				30-34	• •	700		147.5	47.97		40,27
				35+		2 C O	47 11	245.26	73.92		174.39
			D	17-20	1	OLV Lin		181.17	\$ \$3.96		126.09
				21-24	- 1	1.20	12.00	150.66	64.32		83.33
		•		25-29	- 1	13 4 0	70.00	106 75	58.92	10 C	60 . 37
				30-34	• 1	000	20.01	142.59	12.79		65.12
				35+	.0	70V	ev. 90	7-7 A C/C			

For example TABLE 10.1 shows the office premium col (5) applicable on the 1 October 1981 for the following sets of assumptions:

TPBI :	1972-76			
Inflation :	LOW :	AD, TPPD, M 7%;	TPBI	13%
Expense	Fixed at	45% of premium.		

In Chapter 12, the effects of different sets of assumptions on office premium rates are discussed.

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CHAPTER ELEVEN

PRESENTATION OF RESULTS

11.1 Introduction

The traditional way of presenting rates is via a rate book, where every combination of rating factors is defined with its respective rates. In the early 1970's, the Co-operative Insurance Company introduced a 'points table' and several companies are now using similar systems. In the following sections, a method of deriving a premium for a 'points table' is developed. First, a 'points table' is defined. Then follows a discussion why a Company would want to determine its premiums in this way. Then the mathematical background is outlined, and an algorithm for deriving a 'points table' is listed. Finally, an example of the algorithm is discussed.

However, it is emphasised that the 'points table' is only an alternative to the rate book and is sometimes used as a selling point. In addition many companies still use the traditional rating book, however the premiums are based on a 'points table'.

11.2 What is a points table?

The workings of a 'points table' is probably best explained by way of an example. Consider the following:-

TABLE 11.1 - POINTS TABLE

	COMP	NON-COMP			
Cover	19	0			
Policyholder's	17-20	21-24	25-29	30-34	35+
Age	30	12	6	2	0
				•	
	0-3	4-7	8+		
Car Age	8	4 .	0		
Vohiolo	Λ	g	<u> </u>		•
Croup	<u> </u>	<u>ل</u> ح	10	20	
Group	0				

There are four rating factors each with its associated scale. For each point on the scale there is a value, expressed in points.

The procedure used is to record the points score for each rating factor and then aggregate them. Then use a points conversion table to arrive at the premium to be charged. The conversion table is simply a list of points scores with associated monetary values see TABLE 11.2. For example, consider Comprehensive, Policyholder Age 35+, Car Age 4-7 and Vehicle Group A -19 + 0 + 4 + 0 pts. = 23 pts: equivalent to £12.35 from TABLE 11.2.

11.3 Why use a points table?

The workings of a 'points table' are mainly practical and are listed below, not necessarily in order of importance:-

- It is cheaper to produce than the rate book.

TABLE 11.2

POINTS CONVERSION TABLE

Points	Premium	Points	Premium	Points	Premium
	£		£		£
0	5.90	29	14.95	58	37.80
		30	15.45	59	39.00
1	6.10			60	40.25
2	6.30	31	15.95		
3	6.50	32	16.45	61	41.60 .
4	6.70	33	17.00	62	42.95
5	6.95	34	17.55	63	44.35
		35	18.10	64	45.75
6	7.15			65	47.25
7	7.40	36	18.70		
8	7.65	37	19.30	66	48.80
9	7.90	38	19.95	67	50.40
·· 10	8.15	39	20.55	68	52.00
		40	21.25 .	69	53.70
11	8.40			70	55.45
12	8.70	41	21.95		
13	8.95	42	22.65	71	57.25
14	9.25	43	23.40	72	59.10
15	9.55	44	24.15	73	61.05
	`	45	24.95	74	63.00
16	9.85			7 5	65.05
17	10.20	46	25.75		
18	10.50	47	26.55	76	67.20
19	10.85	48	27.45	77	69.35
20	11.20	49	28.35	78	71.60
		50	29.25	79	73.95
21	11.55			80	76.35
22	11.95	51	30.20		
23	12.35	52	31.20	81	78.85
24	12.75	53	32.20	82	81.40
25	13.15	54	33.25	83	84.05
29		55	34.30	84	86.75
26	13.60			85	89.60
27	14.00	56	35.45		
28	14.45	57	36.60	86	92.50

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- It is easy to revise rates.
- The calculation is relatively straightforward.
- The method is easy to understand.

There are however problems associated with such a system. Some are listed below:-

- It may be regarded by brokers and underwriters as over-simplified.
- Fine tuning of the rate structure is no longer possible for marketing purposes. Since in the traditional rating book, particular rates could be altered without affecting any other rates. In a 'point table' this is not possible.
- The points values (as opposed to the monetary values) change only rarely and changes over time in the underlying variables tend to be neglected.
- Since the 'points table' is easy to interpret, any adjustments to the points which reflect statistical experience, are thought by brokers and underwriters to be errors.
- The resulting premium structure is not exact because the underlying analysis is ultimately a simplification.

11.4 THEORY

The problem is to convert the office premium into a 'points table'. Then, it is necessary to compare fitted premium (estimated from our model) with the office premium. That is

for each combination of rating factors, where

∧ OP = fitted premium OP = office premium

Proceed as follows. Let

Then define

where

 $Y = T + CO_i + PH_j + VG_k + CA_l + (In) + e_{ijkl} ...(11.2)$ where the definitions are similar to Section 7.3.2, equation (7.5).

The choice of the figure 1.0325 is quite arbitrary.

In equation (11.1), take logs to the base 1.0325, whence

This relationship is fitted by weighted least squares:-

$$\min \sum_{ijklijkl}^{W} (\log OP_{ijkl} - Y_{ijkl})^2$$
(11.3)

where

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e denotes the least squares estimator

 W_{ijkl} is some set of weights to be selected and CO_i PH_j VG_k and CA_l etc. are the estimated points for the table.

The method of Johnson and Hey (1971) or GLIM can be used to arrive at a 'points table' minimising equation (11.3). However, for simple assumptions, OWLS can be used.

OWLS relies on the fact that it is possible to factorise the weights in equation (11.3), that is

$$W_{ijkl} \propto (co)_{i} (ph)_{j} (vg)_{k} (ca)_{l}$$
 (11.4)

This is demonstrated in Appendix 1.

11.5 An Algorithm

An algorithm follows for the computation of a 'points table'.

- (1) The first step is to choose a set of suitable weights. A suitable choice would be standing business, since, as will later be demonstrated, the method is reasonably robust to the choice of weights selected.
- (2) A standard analysis of variance is performed, using

log W_{ijkl} = log (of standing business) _{ijkl}

- (3) The factors $(i)_{i}$ $(ph)_{j}$ $(vg)_{k}$ $(ca)_{l}$ are selected, after checking for interactions from the analysis of variance.
- (4) The W_{ijkl} is replaced by the product of the estimates of co, ph, vg, ca from the previous step, i.e. then find Y_{ijkl} from the expression

 $\min_{\substack{i \neq l \\ i \neq k}} \sum_{\substack{(co)_{i} \\ (ph)_{j}}} \sum_{\substack{(ca)_{k} \\ (vg)_{l}}} \sum_{\substack{(log \ D^{P}_{ijkl} - Y_{ijkl})}^{2} \cdots (ll.5)$

- (5) The weighted analysis of variance of log OP is then checked to ensure that there are no significant interactions.
- (6) Assuming that no significant interactions appear, then the estimates of the main effects serve as the points within the table.

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- (7) The fitted values for premiums computed through the model are compared with the office premiums, in order to assess the goodness of fit.
- (8) Then an overall adjustment is made to make sure that the sum of the office premium is equal to the sum of fitted premiums.

11.6 Results

11.6.1 Step 1

An example will now be presented. Table 11.8 shows the basic data. The rating factors can be seen on the left. (Column (1) shows the Standard Business, which is being used for the weights, i.e. W_{iikl}.

11.6.2 Step 2

Analysis of variance on log W ijkl gives the following results:-

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TABLE 11.3 - ANALYSIS OF VARIANCE OF LOG W

Factor	Sum of Squares	Degrees of Freedom	<u>Mean Square Ratio</u>
Cover	1.5	1	258
Car Age	5.9	2	487
Vehicle Group	4.6	3	253
Policyholder's Age	33.0	4	1360
Cover x Car Age	12.4	2	1028
Cover x Vehicle Group	0.0	3	1
Cover x PH Age	3.9	4	161
Car Age x Vehicle Group	2.3	6	62
Car Age x PH Age	0.2	8	• 3
Vehicle Group x PH Age	0.5	12	7
Residual	0.5	74	
TOTAL	64.8	119	

From this, it can be seen that, in addition to the main effects, there is a significant association between cover and car age. This point will be returned to later (Section 12.2).

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11.6.3 <u>Step 3</u>

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The standing business distribution yields weights for the factors:-

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TABLE 11.4 - WEIGHTS FOR STANDING BUSINESS

Cover				
<u>Comp</u> . 4.8	Non-Comp. 2.8			
Car Age				
<u>0.3</u> 1.8	<u>4.7</u> 5.3	<u>8+</u> 5•3		
<u>Vehicle (</u>	Group			
A 2.3	B 6.1	<u>C</u> 5•5	D 2.4	
Policyho	lder Age			
<u>17-20</u> 0.9	<u>21-24</u> 1.9	<u>25-20</u> 2.8	<u>30-34</u> 4.3	<u>35+</u> 32•3
		- 104 -		

The values represent the distribution of standing business within the portfolio. For example, there are 32 times as many policies in Age 35+ as in Age 17-20.

11.6.4 <u>Step 4</u>

Using these estimates W ijkl, Y can be estimated from equation (11.5).

11.6.5 <u>Step 5</u>

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The analysis of variance on log OP ijkl yields:-

		•		
Factor	Sum of <u>Squares</u> 1000	Degrees of Freedom	Mean Squares 1000	Mean Square Ratio
	(1)	(2)	(3)=(1)+(2)	(4)=(3)/ Residual(3)
Cover	4351	1	4351	7515
Car Age	3178	2	1589	2745
Vehicle Group	3931	3	1310	2262
Policyholder's Age	2746	4	687	1185
Cover x Car Age	226	2	113	195
Cover x Vehicle Gp.	81	3	27	47
Cover x PH Age	43	4	11	18
Car Age x Vehicle Gp.	. 5	6	1.0	2
Car Age x PH Age	5	8	0.7	1
Vehicle Gp. x PH Age	21	12	2.7	3
Residual	43	74	•6	
TOTAL	14632	119		

TABLE 11.5 - ANALYSIS OF VARIANCE LOG OP

It is important to note that this analysis does not allow us to make any statement about the goodness of fit of the model, since there is not an independent estimate of the residual variance. All that the analysis shows is that, relatively, the main effects are far more important than the interaction terms, which can therefore be neglected.

11.6.6 Step 6

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From these results, we estimate the main effects as follows:-

Weighted Mean	134.1
Cover	
Comp. Non-Comp.	6.3 -10.7
Policyholder Age	
17-20 21-24 25.29 30-34 35+	27.6 20.6 6.8 -2.5 -2.2
<u>Car Age</u>	
0.3 4.7 8+	10.9 4.0 -7.7
Vehicle Group	
Á B C	-8.9 -4.2 0.8

17.0

TABLE 11.6 - ESTIMATES OF POINTS FOR OFFICE PREMIUM

It would be easier to make comparisons if there were no negative values. This is simply accomplished, by transforming the smallest value in each line to zero, making the same addition to the other values in the line, and adjusting the overall mean subsequently.

For example, with car age, we can alter the value for $^{8+'}$ can be altered to zero by adding 7.7, whence 18.6 for $^{0-3'}$ and 11.7 for $^{4-7'}$. The overall mean is later reduced by 7.7, as in TABLE 11.7.

FABLE	11	.7	 ALTERING	THE	BASE	ΤO	FACILITATE	COMPARISON

	Unadjusted	Adjusted
<u>Overall</u>	134.1	104.3
Cover		
Comp. Non-Comp.	6.3 -10.7	17.0 0
Policyholder Age		
17-20 21-24 25-29 30-34 35+	27.6 20.6 6.8 -2.5 -2.2	•30.1 23.1 9.3 .0 0.3
Car Age		
0-3 4-7 8+	10.9 4.0 -7.7	18.6 11.7 .0
Vehicle Group		
A B C D	-8.9 -4.2 0.8 17.0	.0 4.7 9.7 25.9

11.6.7 Step 7

The goodness of fit of the new premium structure can now be examined. There are three possible tests:-

- (1) The fitted premiums can be compared with the office premiums for each rating factor (i.e. Expected/Actual).
- (2) Using 'proposed to existing' method i.e. comparing premium incomes.
(3) The expected premium income from the fitted premiums can also be compared with that of the present premium structure, using the 'proposed to existing' method.

Column (6) of TABLE 11.8 shows the relationship between the fitted premiums and the office premiums, expressed as a percentage. It can be seen that, generally, the divergence is only four percentage points, though in a few cases, e.g. Comprehensive cover for an eight-year-old vehicle in group D driven by a teenager, it is very large. Note also that for Comprehensive vehicle age 0-7, the ratio is in general less than 100, whilst for the corresponding non-comprehensive values, it exceeds 100. This suggests that comprehensive and non-comprehensive require separate tables, as discussed below.

Column (7) of TABLE 11.8 shows the difference between the fitted premiums and office premiums, weighted for the standing business. Column (7) reveals that the actual difference by cell in required premium income is in some cases considerable. It would seem that comprehensive is being undercharged and non-comprehensive overcharged. This will be returned to in section 12.2. However, overall the total premium income is reasonable, as shown by the following:-

(A) Total Standing Business90,362(B) Points Premiums (Fitted using main effects)£8,165,000(C) Office Premiums£8,545,000(D) Difference (B)-(C)-£380,000(E) Ratio (B)/(C) (%)96

OFFICE PREMIUM RESULTS COMPARISON OF ACTUAL V. FITTED

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G.B. FRIVATE CAR PREMIUM ANALYSIS

gas Veh. ACE Veh. ACE OP.H. ACE FAM DE IS STATUTION (1) C2 OTTICE (2) DIFF (3) DIFF (4) DIFF (4) <thdiff (4) DIFF (4) <t< th=""><th></th><th></th><th>E</th><th>XPE</th><th>ENS</th><th>SNE (</th><th>RĂ-</th><th>110 4:</th><th>5% -</th><th>(PD1</th><th>12-76</th><th>LOW</th><th>INFL</th><th>ATION</th></t<></thdiff 			E	XPE	ENS	SNE (RĂ-	110 4:	5% -	(PD1	12-76	LOW	INFL	ATION
$\begin{array}{c c c c c c c c c c c c c c c c c c c $	OVER	VEH. A	AGE	VEH.	GRO	PZH (AGE	STANDING BUSINESS	3 P01N16	FITTED PREMIUM	OFFICE PREMIU	E DIFF	RATIO	SB ×
$ \begin{array}{c c c c c c c c c c c c c c c c c c c $								(1)	(2)	(3)	、 (4)	(5) (3)-(4)	(6) (3)/(4)	(7) (1)≋(5)
$\begin{array}{c c c c c c c c c c c c c c c c c c c $:OMP	· 0-3		A		17-20	>	32	170	270 10				
$\begin{array}{c c c c c c c c c c c c c c c c c c c $						21-24	7	62	163	104.10	244.00	14.73	93	471.43
$\begin{array}{c c c c c c c c c c c c c c c c c c c $						25-29	9	103	142	110.50	131.63	-13.13	21 90	-1351 97
$\begin{array}{c c c c c c c c c c c c c c c c c c c $						30-34	1	152	140	07.92	101.36	-13.44	Úě	-2043.48
$ \begin{array}{c c c c c c c c c c c c c c c c c c c $. '		-		33+		1446	140	80.61	96.89	-8.19	91	-11836.68
$\begin{array}{c c c c c c c c c c c c c c c c c c c $				В		17-20	5 5	52 146	175 170	267.40	279.94	-12.46	95	-648.10
$\begin{array}{c c c c c c c c c c c c c c c c c c c $		•				25-23	7	207	15.4	137.75	200.00 197.91	-19.26	- 91	-2312.65
$\begin{array}{c c c c c c c c c c c c c c c c c c c $		•	•	•	•	30-34	ŧ	472	145	102.20	113.64	16.44	86	7757.51
$ \begin{array}{c c c c c c c c c c c c c c c c c c c $	· •		-		, ·	35+		5063	145	103.00	111.57	8.57	92	-43390.91
$\begin{array}{c c c c c c c c c c c c c c c c c c c $				Ċ		17-20	>	60	100	313.55	324.30	-10.75	76	-644.92
$ \begin{array}{c ccccccccccccccccccccccccccccccccccc$	•					21-24	h D	183	173	250.97	271.73	-20.76	192	-3902.07
$\begin{array}{c c c c c c c c c c c c c c c c c c c $		•				30-34	• . • .	820	157		101.69	-20.21	88	-8974,85
$ \begin{array}{ c c c c c c c c c c c c c c c c c c c$						35+		8142	150	120.75	143.03	-13.42	89 89	-19061.82 -109238.11
$\begin{array}{c c c c c c c c c c c c c c c c c c c $				D	• -	17-20	>	21	196	527.02	487.54	39.40	103	822.05
$\begin{array}{c c c c c c c c c c c c c c c c c c c $					•	21-24	ł	153	139	421.84	420.65	1.19	100	182.19
$\begin{array}{c c c c c c c c c c c c c c c c c c c $			· · .			25-29	7	291	175	271.42	295.06	-23.64	91	6879.41
$ \begin{array}{c c c c c c c c c c c c c c c c c c c $						30-34		630 4490	166	201.37	239.73	-38.36	83	24167.99
$ \begin{array}{c c c c c c c c c c c c c c c c c c c $,		00.		7470	100	202.70	ఉన4.వళ	-21.27	90	-95686.62
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$		4-7	· •	A		17-20		. 35	163	184.74	105.10	.36	22	-12.62
$\begin{array}{c c c c c c c c c c c c c c c c c c c $						21-24	, }	73	136	147.87	152.43	-4.55	97	-423.38
$\begin{array}{c c c c c c c c c c c c c c c c c c c $				•		30-34	•	181	133	70.59	72.66	-2.07	97	-375.10
$ \begin{array}{c c c c c c c c c c c c c c c c c c c $. •			35+		2081	133	71.14	69.37	1.77	102	3691.94
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$	×			В		17-20		94	168	214.75	216.20	-1.45	99	136.29
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$						21-24		229	161	171.09	177.09	-5.20	97	~1109.99
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$,			• · ·		25-27		308 744	147	110.60		1.75	70 94	-679.81
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$						35+		7911	138	82.70	S1.03	1.67	1.2	13241.29
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$				C		17-20		65	173	251.75	242.23	1.02	100	118.26
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$		•• `				21-24		199	166	201.50	207.54	6.04	91	-1201.41
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$						25-29)	437	152	122.65	135.06	-5.41	25	2364.95
$ \begin{array}{c ccccccccccccccccccccccccccccccccccc$					·	30-34 35+	•	801 6877	143	96.19 96.95	27.94		92 90	-6313.66
$\begin{array}{c c c c c c c c c c c c c c c c c c c $				· n ·		17		21	199	423 14	- 	40.18	. 110	847.80
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$				U		21-24		. 113	182	338.69	322.32	9.37	102	1059.02
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$		•				25-29		215	168	217.92	225.30	-7.39	96	-1586.78
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$			< ·			30-34		427	159	161.68	101.11	-19.44	09 07	-8298.80
$\begin{array}{c c c c c c c c c c c c c c c c c c c $			•			35+	· · ·	2818	157	102.70	107.72	-6.77	70	-19020100
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$. 9+		Ą		17-20		17	151	126.94	117.05	2.10	107	154.65
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$		4			• •	21-24		39	144	101.61	99.71	1.90	101	73.91
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$						25-29	•	76	131	65.30	61.36	4.02	106	305.30
$ \begin{array}{c ccccccccccccccccccccccccccccccccccc$				· · · ·		30-34 35+		170	121	43.50 AC 20	45.06 A7 99	- 3.45 A 90	107	506.35 17840-15
$ \begin{array}{c ccccccccccccccccccccccccccccccccccc$								2020		40.07	-0.77			10040.10
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$				В		17-20		43	156	147.57	146.05	1.52	101	65.25
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$						21-24		219	175	70.00	72.29	2.02	102	442.37
$\begin{array}{c c c c c c c c c c c c c c c c c c c $						30-34		437	126	56.38	53.96	2.43	104	1102.90
$ \begin{array}{c ccccccccccccccccccccccccccccccccccc$						35+		6992	125	ರಿತಿ. ಚಿತ	52.23	4.60	100	32161.91
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$				С		17-20		27	161	172.99	163.13	4.85	102	131.07
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$						21-24		60	154	120.46	107.05	1.02	(,'')	- 62.54
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$						20-29		158	140	おんで た	87.62	54	ソワ	-84.61
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$						35+		2346	131	60.1V	64.60	2.01	103	6731.67
$\begin{array}{cccccccccccccccccccccccccccccccccccc$				n		17-20		3	177	299.76	147.36	143.40	197	430.20
25-27 61 157 147.74 155.76 55.75 96 -363.25 30-34 126 147 111.10 124.47 -13.37 59 -1674.55 35+ 966 140 111.97 118.49 55.52 94 5295.66				~		21-24		19	170	232.73	214.45	10.67	103	244.02
$\begin{array}{cccccccccccccccccccccccccccccccccccc$						25-27		61	157	147.74	155.70	5.75	26	-363.25
						ऽQ~-\$4 र5≠		120	147	111,10	124.47	13.37	1314 1214	-14-34.53
						а ст . 		200 	140	****//	110.47	·ひ・ひど	+	

Where col. (1) = standing business at 1 Oct. 80 (3) = Fitted Premium (2) = points applicable to Fitted Prem. (4) = Office Premium

1

TABLE 11.0

OFFICE PREMIUM RESULTS COMPARISON OF ACTUAL V. FITTED

.

G.B. PRIVATE CAR PREMIUM ANALYSIS

EXPENSE RATIO 45% TPB1 72-76 LOW INFLATION

OVER	VEH.	AGE	VEH.	GRO	HZH AGE	STANDING BUGINESS (1)	F01N15	FITTED PREMIUM	OFFICE PREMIUM	01FF (5)	RAT10	SD ≽ D117F (7)
					•				• • •	(3)-(4)	(3)/(4)	(1)*(5)
ION-COMP	0-3		A	*	17-20	6	153	133.42	115.69	17.74		106 42
					21-24	6	146	106.30	77.10	29.69	1.33	178.15
					25-29	6	132	60.71	29.07	39.99	200	233.00
	•				30-34	12	123	50.98	29.60	21.33	172	234.52
					30+	52	123	51.30	47.00	4.30	109	227 .77
			В	i	17-20	15	158	155.10	139.13	15.97	111	239.48
		-			21-24	29	151	124.14	105.53	10.62	117	539.90
					30-34	22	1.002	79.00 59.07	65.02	14.85	122	445.64
					35+	181	128	59.73	43.42 54.63	15.04	100	922.34
		•	с		17-20	17	163	1511 510	15.0 20	177 1 A	-	707 70
					21-24	34	156	145.53	116.97	23.14	1.24	971 10
			•		25-29	- 37	142	93.64	74.40	19.15	125	708.62
			•		30-34	42	133	69.47	52.17	17.30	133	726.45
,					35+	291	133	70.02	6 0.44	9.58	115	2787.07
			D)	17-20	12	179	205.40	224.71	80.82	135	970 .72
	٠.	·			21-24	25	172	244.61 -	177.62	66.29	137	1674.68
		•			25-29	23	158	157.38	113.31	44.07	138	1013.60
		•			30-34	22	149	116.77	3.40	33.37	140	734.09
					35+	121	149	117.68	95.65	22.03	123	2666.02
	4-7	•	А		17-20	85	146	107.12	125.38	-18.25	85	-1551.31
				•	21-24	78	139	85.75	07.93	-2.24	97	174.47
			•		25-29	77	125	55.17	56.11	94	98 1 80	-72.39
1 		د و منتقور ی ب ودند.	- 4 6 		· 30-34,		116	40.73	39.92	1.01	102	100.94
			· ·		35+	482	116	41.25	44.22	2.96	93	-1428.01
			P		17-20	147	151	124.53	142106	-17.53	87	-2576.82
		•	-		21-24	194	144	99.67	99.59	.03	100	10.04
	•		•		25-29	217	1.30	64.13	64.37	24 0 01	104	-01.22 647 17
					30-34 35+	1723	121	47.95	49.54	-1.59	96	- 2733.44
		-		•	47 00	1.00	154	145 99	154 06	8, 09	24	-1132.11
			C	2	17-20	140	108	114.04	103-35	7.99	107	1758.53
					21-24	253	135	75.18	70.91	4.27	106	1080.07
					30-34	300	126	55.73	50.17	5.61	111	1602.52
		••	· · ·	•	35+	1636	126	56.22	54.53	1.69	103	2757.69
			-		17-20	67	172	245.36	224.03	21.34	109	1429.50
		•	· L)	21-24	122	165	196.39	164.92	31.40	119	3840.27
					25-29	130	151	126.36	110.66	15.70	114	2041.08
			•		30-34	174	142	93.75	01.39	12.36	113	2101.0%
					35+	695	142	94.49	04.04	0.40	TON	
	-8+		· •	1	17-20	346	134	73.61	20.60	-25.07	74	-8673.84
					21-24	303	127	50.92	67.93	9.00	30 202	2746.10
					25-29	392	114	37.91	43.81			- 112010140
	•		• •		30-34 35+	504 3217	104 105	20.10	32.78	-4.43	85	-14252.56
6	•							C.E. E. 71	100 07	-24 40	77	-12282.09
			F	3	17-20	528	139	85.57	· 109.97 ウォーロマ	- <u>2</u> 4.4V	90	
					21-24	609	102 115	44.07	40.27	-4.20	21	-3250.07
				· .	25-29	774 	103	32.69	33.27	57	93	-567.51
				•	30-34 35+	6036	107	32.95	36.84	-3.80	ଌ୨	23449.80
				·	17-20	202	144	100.31	119.56	-19.26	83	- 4024.43
			C	2	21-24	234	1.7	30.29	03.07	- 2.73	ンウ	700.94
					25-29	346	123	51.66	53.11	- 1.45	97	-500.57
					30-34	558	114	30.33	37.14	1.19	103	662.43
					35+	2835	114	30.63	40.27	-1.64	95	~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~
			r	0	17-20	56	160	160.60	174.39	5.30	90	- 324.57
			-		21-24	136	153	104.95	126.02	8.06	107	1200.14 1.07 77
					25-27	151	140	86.83	03.33	5.50	104	
					30-34	167	130	64.42	60.CF 25.10	S. 00 19	- <u> </u>	164.40
					35+	808	100	04./U	○↩▖ል⋍	• • /	• •	
· · · · · · · · · · · · · · ·							• •					

Where col. (1) = standing business at 1. Oct 1980 (2) = points applicable to Fitted Premium
(3) = Fitted Premium
(4) = Office Premium • .

Hence the fitted premium would have to be increased by 4% to match the office premium.

A similar comparison may be made with the present premium structure (Table 12.1 column (6)).

(A) Total Standing Business 90,362
 (B) Points Premiums (Fitted using main effects) £8,165,000
 (C) Actual Premiums (charged in the existing £6,538,000 rate book)
 (D) Difference (B)-(C) £1,627,000
 (E) Ratio (B)/(C) (%) 125%

This indicates that a premium increase of about 25% is required to break even (see Section 12.2).

11.6.8 PROBLEMS

There are a number of problems which such a method might encounter in practice, such as:-

Υ.

- (1) The effect on the results if the weights alter?(Section 12.2)
- (2) What happens if we later change the assumptions, e.g. as to expenses or the impact of inflation on claim costs? (Section 12.2)
- (3) If there is association between the main effects at the first stage? (Section 12.2)

These problems are discussed in the next chapter.

CHAPTER TWELVE

COMPARISON OF DIFFERENT SETS OF ASSUMPTIONS

12.1 Comparisons

One of the great advantages of this method is the relative ease with which calculations may be performed. Hence the 'points table' can help examine the effect various input assumptions have on the premium structure. The following changes will be investigated:-

- (a) Inflation: Two scenarios will be tested High, with inflation at 11% for material damage and 17% for liability settlements, and Low, with inflation at 7% for material damage and 13% for liability settlements.
- (b) Base periods: Third Party Bodily Injury claims based on the period 1972-1976 will be compared with data based on the period 1972-1977.
- (c) Expenses: Treating all expenses as fixed, will be compared with the effect of distinguishing between fixed and variable expenses.
- (d) The effect of altering the weights.
- (e) For the main assumptions, comparison with the existing points.
- (f) Separate 'point tables' for comprehensive and non comprehensive.

12.2 Results

A summary of results are shown in TABLE 12.1:-

TABLE	12.	1
		_

Assumptions	(1)	(2)	(3)	(4)	(5)	(6)	
Inflation	Low	High	Low	Low	Low	Present	
TPBI	72-76	72-76	72-77	72-76	72-76	Table	
Expenses	Fixed	Fixed	Fixed	Fixed & Var.	Fixed	•	
Weights	Present	Present	Present	Present	Reversed		
Overall	104	106	100	109	104	104	
Cover							
Comp. Non-Comp.	17 0	16 0	19 0	15 0	19 0	18 0	
Policyholder	Age						
17-20 21-24 25-29 30-34 35+	30 23 9 0 0	30 23 9 0	33 23 10 0 1	27 20 8 0 0	32 23 9 0 2	29 23 8 0 0	
Car Age							
0-3 4-7 8+	19 12 0	19 12 0	19 12 0	16 10 0	13 10 0	7 4 0	
Vehicle Group							
A B C D	0 5 10 26	0 5 10 26	0 5 10 26	0 4 8 23	0 5 9 21	0 5 12 23	

The effect of varying the inflation assumption can be seen by comparing Columns (1) and (2). As expected, the overall level has risen, but the relativities are virtually unchanged.

<u>.</u>..

Changes in the base year assumptions for TPBI are reflected in Columns (1) and (3). The lower overall level suggests that the 72-77 basis yields a lower overall average. This time there has been a change in the policyholder age relativities. The factors affected are cover and policyholder age, reflecting the bases for TPBI input values (Section 8.3.3).

Altering the expense assumptions from Fixed (Column (1)) to Fixed plus Variable (Column (4)), gives results as expected, namely, an increase in overall level combined with a narrowing of the relativities for the fixed and variable basis.

To investigate the effect of changing the weights, an experiment was undertaken. The weights used in the previous example - Column (1) of TABLE 11.8 - form a series, starting 32, 62, 103... and ending ...151, 167, 858. This series was reversed, i.e. the weights used on the second occasion began 858, 167, 151... and ended 103, 62, 32.

Column (5) shows the results obtained when the weights were reversed. It can be seen that the overall level does not change but the relativities do. Given the drastic nature of the change, the fitting procedure seems reasonably robust to portfolio changes.

Comparing column (6) which represents the existing premium table, with say column (1) the major difference lies in the car age. Namely that newer cars are being significantly undercharged. Any changes to the existing table will depend on market considerations.

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In Section 11.6.2 brief mention was made of the question of what is to be done if associations between main effects are observed at the first stage of the OWLS analysis. In TABLE 11.3, an association was noticed between cover and age of car. One possible solution is to have separate rating structures for the two different types of cover. The results are summarised in TABLE 12.2:-

	Comprehensive	Non-Comprehensive
Overall	118	105
Policyholder Age		•
17-20 21-24 25-29 30-34 35+	28 23 10 2 0	36 25 11 0 3
Car Age		
0-3 4-7 • 8+	23 13 0	11 9 0
Vehicle Group	•	
A B C D	0 5 11 28	0 4 7 21

TABLE 12.2 - SEPARATE TABLES FOR COMP AND NON COMP

These results suggest that there may be valid theoretical grounds for having separate premium structures. In particular notice that car age and vehicle group are less important under the non comprehensive points. This seems reasonable, since there is no AD cover. As decisions are taken on practical as well as theoretical grounds and it is highly probable that despite the above results, a single premium structure for comprehensive and non comprehensive will be selected, because of marketing considerations.

This now concludes the technical analysis of the premium bases.

CHAPTER THIRTEEN

MARKETING ASPECTS

The Group will now be in a position to discuss the premium recommendation.

Factors they will have to consider will include inter alia:-

- (i) what other companies have done since the last meeting;
- (ii) whether the competitive position will allow all or some of the recommendations to be implemented. For example, in TABLE 12.1, compare the actuarial premium recommendations Column (1) with the existing structure Column (6). The outstanding differences occur when comparing the points for car age. Hence the actuarial recommendation might indicate an increase in rates for new cars. However, this is unlikely to be acceptable because of the market conditions, which give little weight to this rating factor;
- (iii) whether there are to be rate changes, their timing, and the likely reaction of the market;
- (iv) any special marketing campaign proposed, e.g. introducing new factors, offering discounts for older drivers, or introducing protected no claim discount.

As a final thought, the market place for selling motor insurance is changing, owing to the advent of direct telecommunication on television. The time will come (when the next generation of televisions are made) when quotations will be available to the public via the television in their own homes. Insurance companies and brokers will have to consider the implications of this new dimension. This just highlights the dynamic world of marketing. This is already being used by the Automobile Association, who have many companies quoted in their scheme.

The issues raised by marketing considerations are very important, and in fact dominate any decision process. However, the subject is outside the scope of this thesis.

ANALYSIS OF SURPLUS

14.1 Theory

Premium rates are simply an attempt to forecast claim and expense experience. For proper control an analysis of surplus should be regularly performed to measure where the forecasts are failing and the effects on the profitability of business. Lee (1973) gives an example which is the basis of this work.

The following analysis of surplus for motor is based on a note by Grant (1981). Additional papers by Brennan (1968) and Taylor (1974) are also relevant. These latter papers discuss the problem of ordering of the analysis. This is similar to the problem encountered in the analysis of variance for the unbalanced design. Also the same problem is found in Pensions fund analysis of surplus when salary and withdrawal are ordered Lee (1973).

In 1977, a projection based on data then available was made for 1978; the main categories of data are listed below. The results actually experienced differ from those projected because the assumptions made in the projection do not wholly agree with the actual experience. What is needed is a method of analysing these differences into their component contributions which can later be added back to yield the differences between actual and projected results. This problem is similar to an analysis of surplus and the information gained will be useful in fine-tuning the assumptions in the model, as the monetary effect of the differences between the assumptions made in the model and the actual experience is discovered. It would also provide a useful format for analysing these differences for management.

Companies use models of which the simplest involve only premiums and claim ratios, while the more complex ones involve assumptions about exposure, average premium rate, average claim cost, claim frequency and expenses. In the model considered, assumptions on the following factors are made for each period and risk group:-

- SB Standing Business (number of vehicles at the end of the year)
- AWP Average Written Premium
- ACF Average Claim Frequency
- ACC Average Claims Cost
- VER Variable Expense Rates (this would include the commission rate)
- FE Fixed Expenses

The Earned Premium EP and the Exposure EX can be easily calculated using the ¹/8 method for quarterly projections or the ¹/24 method for monthly projections (Benjamin (1977)). The Average Earned Premium AEP can be calculated by dividing the EP by the EX. (At present, tax and investment income are ignored). The result from the year's Underwriting Surplus (US) (i.e. excluding any adjustment to outstanding claims for prior years) can be expressed in the form

US = Earned Premium - Claims - Expenses (inclusive of commission)

The Underwriting Surplus US projected by the model is

US = EX x AEP - EX x ACF x ACC - SB x VER x AWP - FE

Denote by ' the figures derived from the actual results. It is possible to derive the AEP' from the Earned Premium EP' and the Exposure EX'. The claims can easily be expressed in the form EX' x ACF' x ACC' as the Exposure EX' and the Claim Frequency ACF' are known. The Underwriting Surplus can, therefore, be expressed as

US' = EX' x AEP' - EX' x ACF' x ACC' - SB' x VER' x AWP' - FE'

١.

The differences between the actual experience and that projected are the result of differences in the exposure, claim frequency, level of expenses, inflation rate, etc. The formulae for a possible analysis of assessing the numerical contribution from each of the factors are given below.

Effects of

Exposure

 $(EX' - EX) \times (AEP - ACF \times ACC) - (SB' - SB) \times VER \times AWP$

٠

Average Premium

 $EX' \times (AEP' - AEP) - SB' \times VER \times (AWP' - AWP)$

Claims Frequency

- (EX' x (ACF' - ACF) x ACC)

Claims Inflation

- (EX' $x \text{ ACF'} x (\text{ACC}^* - \text{ACC})$)

Claims Cost

- (EX' x ACF' x (ACC' - (ACC*))

.

Expenses

- (FE' - FE + SB' x VER' x AWP' - SB' x VER x AWP')

(where ACC* is the forecast of the average claims cost using either the known or the latest estimates of claims cost inflation rates).

The formulae above can easily be checked from the expression for Underwriting Surplus.

The effect of claim frequency and claims cost and claims inflation can be expressed in several different ways. Two of these are given below:-

Claim	Original	Alternative			
Frequency	- EX' x (ACF' - ACF) x ACC	- EX' x (ACF' - ACF) x ACC'			
Claim Cost	- EX' x ACF' x (ACC' - ACC*)	- EX' x ACF x (ACC' - ACC*)			
Claims Inflation	- EX' x ACF' x (ACC* - ACC)	- EX' x ACF x (ACC* - ACC)			

The Original is preferred and is used in in the analysis because the effect of the claim frequency is independent of the actual claims cost. The actual claims cost wil be partly based on the outstanding claims estimated and will change over time until the ultimate settlement. Any adjustment to these estimates in the analysis will affect only the result attributed to claims cost.

Various other alternative breakdowns of the analysis exist and the choice largely depends on the projection procedure, the data and the factors to be highlighted. For example, Grant (1981) considers, the effect of dividing average claims cost into its component parts, namely: AD, TPPD, TPBI, M and inflation.

14,4 Analysis of Surplus - Example

The following example is taken from the forecasts of part of the Private Car portfolio of an Insurance company. The forecasts are made on a year of accident basis, this example being 1978. Four sets of forecasts for 1978 are given, corresponding to a projection at the end of each quarter in 1977.

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The actual results for claims cost, include, the latest case estimates and therefore this analysis would be subject to some adjustments as the claim payments run off.

	1st	2nd	3rd	4th	
(Figures in '000s)	Forecast	Forecast	Forecast	Forecast	Results
Standing Business	101.4	88.4	87.4	85.4	78.3
Exposure	97.5	87.8	87.8	85.8	80.8
Written Premium	6244.	5702.	6114.	6339.	5794.
Earned Premium	6061.	5684.	5902.	6035.	5600.
No. of Claims	13.1	12.6	12.4	12.2	- 11.7
Claims Costs	3993.	3938.	3821.	3712.	3348.
Expenses +	2077.	1987.	2184.	2295.	2342.
Profit (Loss) *	(9)	(241)	(103)	28	(90)

* Profit = Earned Premium - Claims Cost - Expenses

+ including Variable Expenses (15% of Written Premium)

To calculate the effects for the analysis of surplus we need, •• in addition, to recalculate the forecasts of claims cost using the latest estimates of claims inflation. These averages are given below:-· .

、	1st Forecast	2nd Forecast	3rd Forecast	4th Forecast	Results
Forecasted Average	304.81	312.54	308.15	304.26	286.15
Adjusted Forecast	289.85	300.13	304.13	302.84	

The effect of each factor is then:

	1st	2nd	3rd	4th
(Figures in '000s)	Forecast	Forecast	Forecast	Forecast
Frosure	-141	-41	-71	-56
Premiums	431	258	121	-81
Claim Frequency	-257	-33	-89	-64
Claims Inflation	175	145	47	17
Claims Costs	43	164	210	195
Expenses	-333	-341	-200	-129

The initial underestimation of the average premium is partly due to the premium increases during 1977 and 1978 which were not allowed for in the earlier forecasts.

Initially claim frequency is very important. However, as the forecast approaches 1978, the estimate of frequency improves, however a 4% over prediction in frequency produces -64 contribution to surplus.

The effect of claims cost unexpectedly increases between the first and last forecasts, highlighting either a need to examine the methods used to project claims cost, or possibly an unexpected change in the claims experience.

The relatively large effect of expenses is due to two problems. Firstly, there is a slight difference between forecasts and results in the basis for allocating fixed expenses between classes; also, the inflation rates used to project expenses during 1978 could be out of line with those actually experienced. An obvious extension of these methods would be to separate the effect of inflation from effect of expenses in a similar manner to that employed for claims costs.

CONCLUSION

15.1 Summary

This thesis has put forward a framework for the analysis of motor insurance premium rating, which combines both practical and technical judgements. It has argued for a detailed breakdown of the data. This allows for an in-depth analysis of claim proportions and claims cost. Since the data-set analysed were relatively small, sophisticated statistical modelling techniques were employed to get an understanding of the underlying structure and smooth out variations in the observed data. It is suggested that the results from this part of the thesis encourage similar analyses by insurance companies, and they need not rely on industry statistics to obtain input for a premium basis.

Then by combining claims experience data together with economic views concerning inflation and expenses an office premium was calculated. The advantages and disadvantages of converting the office premium into a 'points table' were briefly discussed. A mathematical method to arrive at the 'points table' was developed and put into a simple algorithm. One of the advantages of a 'points table' is the ease of comparing different sets of assumptions in the premium basis. As an example several sets of assumptions were analysed and discussed. Finally as one of the themes of this thesis is to encourage explicit account of the various assumptions in the premium basis, an analysis of surplus was developed. This type of analysis is important as a learning process since the assumptions can be monitored by comparing actual against expected, future premium analysis will benefit.

15.2 Future Research

Throughout this work comments concerning areas of future research were made. In summary the main areas are;-

- The effect on premium rates, if investment income were explicitly taken into account in the office premium. This would involve looking at net cash flow between premiums and claims, and not necessarily using the simple solution of average settlement dates. Also it involves the allocation of assets between types of business written by the company to arrive at a rate of return.
- Modelling techniques for claims cost, in particular TPBI, are still in their early stages of development.
- Very little work has been published (e.g. (1979)) on how well the models used to fit data perform as forecasting tools. Since premium rating is forecasting, research in this area is very important.

- The worked example assumed that the claims cost inflation rates obtained from the EAG were correct. Some formal analyses are required to judge if these assumptions are in fact reasonable.
- Could an alternative minimising method be found to arrive at a 'points table' which will give a closer fit within the portfolio.

••• -

APPENDIX 1

Orthogonal Weighted least Squares Analysis

The general problem is to minimise the following expression:-

$$\sum_{ijk} \frac{1}{V_{ijk}} \left[\begin{array}{c} Z_{ijk} \left(\hat{S} + \hat{R}_{i} + \hat{U}_{j} + \hat{V}_{k} + (\hat{I}n) \right) \right]^{2} \dots (A.1)$$
where Z_{ijkl} is the dependent variable for cell ijkl
 V_{ijkl} is the weights for cell ijkl
 S is the overall mean

Ri is the effect of R at level i

Uj is the effect of U at level j

is the effect of V at level k Vk

is the effect of the interaction terms (In)

is the least squared estimate

e.g. Johnson and Hey (1971) used n as the weight,

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and Seal (1968) used $\frac{P(1-P)}{n}$ as the weight

Please (1974), suggested that the minimisation could be simplified by assuming that the weight $\frac{1}{V_{ijk}}$ factorises i.e. if V_{ijk} satisfies

$$\frac{1}{\mathbf{V}_{ijk}} d r_i u_j v_k \text{ (for all ijk)} \qquad \dots \text{ (A.2)}$$

and various concomitant constraints. The design becomes orthogonal, enabling the following estimators to be derived: $\underbrace{ \underbrace{ \sum_{i j k} \sum_{k} \left(\begin{array}{cc} \frac{r_{i}}{r_{\cdot}} & \frac{u_{j}}{u_{\cdot}} & \frac{v_{k}}{v_{\cdot}} \right) (Z_{ijk}) }_{\text{II}}$ **S** = $5 = \frac{u_i}{v_k} \frac{v_k}{(7_i + i_k)} - S$ Ri

$$= \sum_{j=1}^{\infty} \sum_{k=1}^{\infty} \frac{1}{2} \frac{1}{2} \frac{1}{k} \left(\frac{1}{2} \frac{1}{jk} \right) - \frac{1}{2}$$

$$(RU)_{ij} = \sum_{k} \frac{V_{k}(Zijk) - S - R_{i} - U_{j}}{V_{i}}$$

and similar equations for $(RV)_{ik}$ and $(UV)_{jk}$ where $r = \sum_{i}^{j} r_{i}$, $u = \sum_{j}^{j} u_{j}$ and $v = \sum_{k}^{j} v_{k}$

and the constraints are

$$\sum_{i} r_{i} R_{i} = \sum_{j} u_{j} U_{j} = \sum_{k} v_{k} V_{k} = 0$$

$$\sum_{i} r_{i} (RU)_{ij} = \sum_{j} u_{j} (RU)_{ij} = 0 \text{ etc.}$$

One method of justifying the factorisation, is by an analysis of $-\log V_{ijk}$. Under the assumption of perfect factorisation

 $-\log V_{ijk} = s' + r'_{i} + u'_{j} + v'_{k}$

where s' is an arbitory constant and r'_i, u'_j and v'_k correspond to r_i , u, and V_k

in equation (A.2). Thus, if the interactions in an analysis of variance of - $\log V_{ijk}$ are negligible, there is evidence supporting the factorisation and hence the main effects can be used to calculate the weights.

In practice the weights do not always factorise perfectly, but experience shows that the model is robust with respect to the weights chosen.

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