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# First-Differenced Inference for Panel Factor Series

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## Abstract

We complement existing inferential theory for panel factor models by deriving the asymptotics for the first differences of the estimated factors and common components obtained from a non-stationary panel factor model. As an application, we propose an estimator for the long run variance of the common components.

**JEL Classification:** C13, C23.

**Keywords:** Non-stationary panels, common factors, common components, first differences.

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# 1 Introduction

Consider the non-stationary panel factor series

$$X_{it} = \lambda'_i F_t + e_{it}, \quad (1)$$

where  $i = 1, \dots, n$ ,  $t = 1, \dots, T$ ,  $F_t$  is a  $k$ -dimensional vector with DGP  $F_t = F_{t-1} + \varepsilon_t$ , and  $e_{it}$  is stationary. Bai (2004) develops the inferential theory for (1) - specifically, for  $F_t$ ,  $\lambda_i$ , and for the non-stationary common component  $C_{it} \equiv \lambda'_i F_t$ . Alternatively, one may also consider the stationary, first-differenced model

$$x_{it} = \lambda'_i f_t + u_{it}, \quad (2)$$

where  $x_{it} = \Delta X_{it}$  and  $f_t = \Delta F_t$ . In this case, estimators for  $\lambda_i$ ,  $f_t$  and  $c_{it} \equiv \lambda'_i f_t$  ( $\hat{\lambda}_i$ ,  $\hat{f}_t$  and  $\hat{c}_{it}$  respectively) are provided by Bai (2003).

This note complements the existing inferential theory on (1) and (2), by studying estimation based on the first difference of the estimator of  $F_t$ , say  $\hat{F}_t$ , computed from (1). Indeed, instead of estimating  $f_t$  from (2), one could use  $\tilde{f}_t = \hat{F}_t - \hat{F}_{t-1}$ . Thence, using either the estimated  $\lambda_i$  from (1), say  $\hat{\lambda}_i$ , or estimating  $\lambda_i$  from (2) using  $\tilde{f}_t$ , one can compute the first differenced estimator of  $c_{it}$  as  $\tilde{c}_{it} \equiv \tilde{\lambda}'_i \tilde{f}_t$ . Estimating  $f_t$  and  $c_{it}$  is useful for various purposes; in this paper we consider the estimation of the long run covariance matrices (henceforth, LRV) of  $F_t$  and  $C_{it}$ .

Some results have already been developed by Trapani (2012) in the context of bootstrapping nonstationary factor models. This note completes the inferential theory for the first-differenced estimators, reporting rates of convergence for:  $\tilde{f}_t$ ; for the estimator of  $\lambda_i$  based on  $\tilde{f}_t$ , say  $\tilde{\lambda}_i$ ; and for a weighted-sum-of-covariances estimator of the LRV of  $C_{it}$  based on  $\tilde{f}_t$ .

## 2 Results

All results are derived under the same assumptions as in Bai (2003, 2004), omitted for brevity. Henceforth, we define the  $r \times r$  rotation matrix  $H \equiv \left( \frac{\hat{F}'F}{T^2} \right) \left( \frac{\Lambda'\Lambda}{n} \right)$ , where  $F = [F_1, \dots, F_T]'$  ( $\hat{F}$  is defined similarly) and  $\Lambda = [\lambda_1, \dots, \lambda_n]'$ . The number of factors,  $r$ , is assumed known.

We firstly report a Lemma containing rates of convergence for  $\tilde{f}_t = \hat{F}_t - \hat{F}_{t-1}$ .

**Lemma 1** *As  $(n, T) \rightarrow \infty$ , it holds that*

$$\tilde{f}_t - H'f_t = O_p\left(\frac{1}{\sqrt{n}}\right) + O_p\left(\frac{1}{T^{3/2}}\right), \quad (3)$$

$$\max_{1 \leq t \leq T} \|\tilde{f}_t - H'f_t\| = O_p\left(\frac{1}{T}\right) + O_p\left(\sqrt{\frac{T}{n}}\right), \quad (4)$$

$$\frac{1}{T} \sum_{t=1}^T (\tilde{f}_t - H'f_t) u_{it} = O_p\left(\frac{1}{\sqrt{n}}\right) + O_p\left(\frac{1}{T^{3/2}}\right). \quad (5)$$

Under  $\frac{n}{T^3} \rightarrow 0$ ,  $\sqrt{n}(\tilde{f}_t - H'f_t) \xrightarrow{d} QN(0, \Upsilon_t)$ , where  $Q$  is defined in Theorem 2 in Bai (2004, p. 148) and  $\Upsilon_t \equiv \lim_{n \rightarrow \infty} n^{-1} \sum_{i=1}^n \sum_{j=1}^n E(\lambda_i \lambda_j' u_{it} u_{jt})$ .

Lemma 1 states that rates and uniform convergence of  $\tilde{f}_t - H'f_t$  are the same as for  $\hat{F}_t - H'F_t$  - see Lemma 2 in Bai (2004). This can also be compared with the results in Theorem 2 in Bai (2003), where it is shown that  $\hat{f}_t - H_1'f_t = O_p(n^{-1/2}) + O_p(T^{-1})$  - in general, the rotation matrices  $H$  and  $H_1$  are different. Therefore, heuristically,  $\tilde{f}_t$  should be a better estimator than  $\hat{f}_t$  for the space spanned by  $f_t$ , especially when  $T$  is small. Lemma 1 is a complement, regarding the properties of  $\tilde{f}_t$ , to Lemma A.1 in Trapani (2012).

We now turn to presenting results on the estimation of the loadings  $\lambda_i$ . To this end, it is possible to use the estimator of  $\lambda_i$  from (1), say  $\hat{\lambda}_i$ . Bai (2004, p. 148-149) shows that  $\hat{\lambda}_i$  is “superconsistent”, viz.  $\hat{\lambda}_i - H^{-1}\lambda_i = O_p(T^{-1})$ ; also, the rate of convergence does

not depend on  $n$ . Alternatively, it is possible to estimate loadings as  $\tilde{\lambda}_i = \left[ \sum_{t=1}^T \tilde{f}_t \tilde{f}_t' \right]^{-1} \left[ \sum_{t=1}^T \tilde{f}_t x_{it} \right]$ . Let  $\Sigma_\varepsilon \equiv E(\varepsilon_t \varepsilon_t') = E(f_t f_t')$ ; it holds that:

**Proposition 1** *As  $(n, T) \rightarrow \infty$  it holds that  $\tilde{\lambda}_i - H^{-1} \lambda_i = O_p(n^{-1}) + O_p(T^{-1/2})$ . Under  $\frac{\sqrt{T}}{n} \rightarrow 0$ ,  $\sqrt{T} \left( \tilde{\lambda}_i - H^{-1} \lambda_i \right) \xrightarrow{d} N(0, V_i)$  with  $V_i = (H' \Sigma_\varepsilon H)^{-1} (H' \Phi_i H) (H \Sigma_\varepsilon H')^{-1}$  and  $\Phi_i = \lim_{T \rightarrow \infty} E(f_t f_s' u_{it} u_{is})$ .*

Proposition 1 states that the properties of  $\tilde{\lambda}_i$  are (apart from the rotation matrix  $H$ ) the same as in Theorem 2 in Bai (2003), where estimation of  $\lambda_i$  is based on using (2). This can be compared with  $\hat{\lambda}_i$ , whose convergence rate does not depend on  $n$  and it is faster in  $T$ .

Based on Lemma 1 and Proposition 1, consider the first-differenced estimator of the common components  $c_{it}$ ,  $\tilde{c}_{it} \equiv \hat{\lambda}_i' \tilde{f}_t = \hat{C}_{it} - \hat{C}_{it-1} = \hat{\lambda}_i' (\hat{F}_t - \hat{F}_{t-1})$ . By combining the results above, and using Lemma 3 in Bai (2004), we have  $\tilde{c}_{it} - c_{it} = \hat{\lambda}_i' \tilde{f}_t - \lambda_i' f_t = \left( \hat{\lambda}_i - H^{-1} \lambda_i \right)' \tilde{f}_t + \left( \tilde{f}_t - H' f_t \right)' H^{-1} \lambda_i + \left( \hat{\lambda}_i - H^{-1} \lambda_i \right)' \left( \tilde{f}_t - H' f_t \right) = O_p(n^{-1/2}) + O_p(T^{-1})$ . Using Theorem 3 in Bai (2004) on the limiting distribution of  $T \left( \hat{\lambda}_i - H^{-1} \lambda_i \right)$ , the asymptotic distribution of  $\tilde{c}_{it} - c_{it}$  has the same properties as in Theorem 4 in Bai (2004, p. 149).

The results in Lemma 1 and Proposition 1 can be combined in order to estimate the LRV of  $F_t$  and  $C_{it}$ . Let  $\Sigma_F$  be the LRV of  $F_t$ , and define similarly the LRV of  $C_{it}$  as  $\Sigma_C$ . A rotation of  $\Sigma_F$  can be estimated as

$$\hat{\Sigma}_F = \hat{\gamma}_0^F + \sum_{j=1}^h \left( 1 - \frac{j}{h+1} \right) (\hat{\gamma}_j^F + \hat{\gamma}_j^{F'}),$$

where  $h$  is a bandwidth parameter and  $\hat{\gamma}_j^F \equiv T^{-1} \sum_{t=j+1}^T \tilde{f}_t \tilde{f}_{t-j}'$ . Of course,  $\hat{\Sigma}_F$  does not estimate  $\Sigma_F$  consistently due to rotational indeterminacy; it can be expected that  $\left\| \hat{\Sigma}_F - H' \Sigma_F H \right\| = o_p(1)$ . Similarly,  $\Sigma_C$  can be estimated either as  $\hat{\Sigma}_C = \hat{\lambda}_i' \hat{\Sigma}_F \hat{\lambda}_i$ , or as

$\tilde{\Sigma}_C = \tilde{\lambda}'_i \hat{\Sigma}_F \tilde{\lambda}_i$ . By virtue of Proposition 1,  $\hat{\Sigma}_C$  should be better, and we focus our attention on it.

**Theorem 1** *Assume that  $\sum_{j=0}^{\infty} j^s |\gamma_j^F| < \infty$ . It holds that*

$$\left\| \hat{\Sigma}_C - \Sigma_C \right\| = O_p \left( \frac{h}{\sqrt{T}} \right) + O_p \left( \frac{h}{n} \right) + O_p \left( \frac{1}{h} \right). \quad (6)$$

Theorem 1 contains rates of convergence for  $\hat{\Sigma}_C$ , which is consistent provided that  $h \rightarrow \infty$  and  $h / \min \{n, \sqrt{T}\} \rightarrow 0$ . This also gives a selection rule for  $h$ ; the choice of the bandwidth that maximizes the speed of convergence is  $h^* = O(\min \{T^{1/4}, n^{1/2}\})$ .

We point out that  $\hat{\Sigma}_C$  is not the only possible estimator for  $\Sigma_C$ . One could consider estimating a rotation of  $\Sigma_F$  using  $\hat{f}_t$  calculated from (2). Given that  $H$  differs depending on whether (1) or (2) is used, in this case it is necessary to employ the estimated loadings from model (2), which have the same properties as  $\tilde{\lambda}_i$  in Proposition 1. Based on this, and on Lemma 1, it can be expected that this estimator does not converge as fast as  $\hat{\Sigma}_C$ . Similarly, it is possible to estimate  $\Sigma_C$  using the  $x_{it}$ s directly. Theoretically, this estimator should work, since  $e_{it}$  is stationary, although this may introduce some noise in the estimation of  $\Sigma_C$ .

## Proofs

**Proof of Lemma 1.** See the online material.

**Proof of Proposition 1.** Let  $\delta_{nT} \equiv \min \{\sqrt{n}, T\}$ . By definition,  $\tilde{\lambda}_i - H^{-1}\lambda_i = \left( \sum_{t=1}^T \tilde{f}_t \tilde{f}'_t \right)^{-1} \times \left[ \sum_{t=1}^T H' f_t u_{it} + \sum_{t=1}^T \tilde{f}'_t \left( \tilde{f}_t - H' f_t \right) \lambda_i + \sum_{t=1}^T \left( \tilde{f}_t - H' f_t \right) u_{it} \right] = \left( \sum_{t=1}^T \tilde{f}_t \tilde{f}'_t \right)^{-1} (I + II + III)$ . Consider the denominator. By Lemma A.1 in Trapani (2012),  $\sum_{t=1}^T \left\| \tilde{f}_t - H' f_t \right\|^2 = O_p(T\delta_{nT}^{-2})$  and  $\sum_{t=1}^T \left( \tilde{f}_t - H' f_t \right)' f_t = O_p(\sqrt{T}\delta_{nT}^{-1}) + O_p\left(\frac{\sqrt{T}}{n}\right)$ . Hence,  $\sum_{t=1}^T \tilde{f}_t \tilde{f}'_t = H' \sum_{t=1}^T f_t f'_t H + o_p(T) = O_p(T)$ . As regards the numerator,  $I = O_p(\sqrt{T})$  by a CLT. Using the same arguments as for the denominator,  $II = O_p(\sqrt{T}\delta_{nT}^{-1}) + O_p\left(\frac{\sqrt{T}}{n}\right)$ . Hence,  $\tilde{\lambda}_i - H^{-1}\lambda_i = O_p(T^{-1/2}) + O_p(n^{-1})$ . Finally,  $III = O_p(n^{-1/2}) + O_p(T^{-3/2})$  using (5). The

limiting distribution follows from noting that, when  $\frac{\sqrt{T}}{n} \rightarrow 0$ , the dominating  $O_p(T^{-1/2})$  term is  $\left(H' \sum_{t=1}^T f_t f_t' H\right)^{-1} \left(\sum_{t=1}^T H' f_t u_{it}\right)$ .

**Proof of Theorem 1.** We omit  $H$  for simplicity when this does not cause ambiguity. We start by showing that  $\left\|\hat{\Sigma}_F - H' \Sigma_F H\right\| = O_p\left(\frac{h}{\sqrt{T}}\right) + O_p\left(\frac{h}{n}\right) + O_p\left(\frac{1}{h}\right)$ . By definition,  $\Sigma_F = \gamma_0^F + \sum_{j=1}^{\infty} (\gamma_j^F + \gamma_j^{F'})$ , whence

$$\begin{aligned} \hat{\Sigma}_F - \Sigma_F &= (\hat{\gamma}_0^F - \gamma_0^F) + \sum_{j=1}^h \left(1 - \frac{j}{h+1}\right) [(\hat{\gamma}_j^F + \hat{\gamma}_j^{F'}) - (\gamma_j^F + \gamma_j^{F'})] \\ &\quad - \sum_{j=1}^h \left(\frac{j}{h+1}\right) (\gamma_j^F + \gamma_j^{F'}) - \sum_{j=h+1}^{\infty} (\gamma_j^F + \gamma_j^{F'}) \\ &= I - II - III. \end{aligned}$$

Consider  $I$ . We have  $\hat{\gamma}_0^F - \gamma_0^F = T^{-1} \sum_{t=j+1}^T \tilde{f}_t \tilde{f}_t' - \gamma_0^F = \left(T^{-1} \sum_{t=j+1}^T f_t f_t' - \gamma_0^F\right) - T^{-1} \sum_{t=j+1}^T (\tilde{f}_t - f_t) f_t' - T^{-1} \sum_{t=j+1}^T f_t (\tilde{f}_t - f_t)' + T^{-1} \sum_{t=j+1}^T (\tilde{f}_t - f_t) (\tilde{f}_t - f_t)' = I_a + I_b + I_b' + I_c$ . The CLT yields  $I_a = O_p(T^{-1/2})$ ; as far as  $I_b$  and  $I_c$  are concerned, Lemma A.1 in Trapani (2012) entails that they are both  $O_p(n^{-1}) + O_p(T^{-2})$ . The same holds for  $\hat{\gamma}_j^F - \gamma_j^F$ ; putting all together  $I = O_p(hT^{-1/2}) + O_p(hn^{-1})$ . Standard arguments yield  $II = O(h^{-1})$  and  $III = o(h^{-s})$ . The Theorem follows from  $\hat{\lambda}_i - H^{-1} \lambda_i = O_p(T^{-1})$ .

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